METHODOLOGY

Using an innovative multiple regression procedure in a cancer population (Part II): fever, depressive affect, and mobility problems clarify an influential symptom pair (pain–fatigue/weakness) and cluster (pain–fatigue/weakness–sleep problems)

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Background: Most patients with advanced cancer experience symptom pairs or clusters among pain, fatigue, and insomnia. However, only combinations where symptoms are mutually influential hold potential for identifying patient subgroups at greater risk, and in some contexts, interventions with "cross-over" (multisymptom) effects. Improved methods to detect and interpret interactions among symptoms, signs, or biomarkers are needed to reveal these influential pairs and clusters. I recently created sequential residual centering (SRC) to reduce multicollinearity in moderated regression, which enhances sensitivity to detect these interactions.

Methods: I applied SRC to moderated regressions of single-item symptoms that interact to predict outcomes from 268 palliative radiation outpatients. I investigated: 1) the hypothesis that the interaction, pain × fatigue/weakness × sleep problems, predicts depressive affect only when fever presents, and 2) an exploratory analysis, when fever is absent, that the interaction, pain × fatigue/weakness × sleep problems × depressive affect, predicts mobility problems. In the fever context, three-way interactions (and derivative terms) of the four symptoms (pain, fatigue/weakness, fever, sleep problems) are tested individually and simultaneously; in the non-fever context, a single four-way interaction (and derivative terms) is tested.

Results: Fever interacts separately with fatigue/weakness and sleep problems; these comoderators each magnify the pain–depressive affect relationship along the upper or full range of pain values. In non-fever contexts, fatigue/weakness, sleep problems, and depressive affect comagnify the relationship between pain and mobility problems.

Conclusion: Different mechanisms contribute to the pain × fatigue/weakness × sleep problems interaction, but all depend on the presence of fever, a sign/biomarker/symptom of proinflammatory sickness behavior. In non-fever contexts, depressive affect is no longer an outcome representing malaise from the physical symptoms of sickness, but becomes a fourth symptom of the interaction. In outpatient subgroups at heightened risk, single interventions could potentially relieve multiple symptoms when fever accompanies sickness malaise and in non-fever contexts with mobility problems. SRC strengthens insights into symptom pairs/clusters.

Keywords: depression, moderated regression, multicollinearity, sickness behavior, statistical interaction, symptom cluster

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Introduction

Patients treated with radiation or chemotherapy often experience pairs or clusters of symptoms that occur either simultaneously, or within the same period (for a recent

review, see Kirkova et al[1]). An assessment of findings from several studies concluded that more than half of all patients receiving treatment for advanced cancer experience pairing or clustering among the symptoms of pain, fatigue, and insomnia.[2]

In Part I of this two-part article, I revealed that in an outpatient sample receiving palliative radiation to relieve painful bone metastases, a three-way symptom interaction of pain, fatigue/weakness, and sleep problems predicted higher levels of depressive affect in moderated regression when these symptoms co-occurred in participants as a pair (pain–fatigue/weakness) or as a cluster (pain–fatigue/weakness–sleep problems).[3] Previous studies have also revealed this symptom pair[4] or cluster[5] in patients receiving chemotherapy or initiating palliative radiotherapy for bone metastases. In these individuals, pairs or clusters of physical symptoms may result from elevated production of proinflammatory cytokines (cell products released during immunological reactions) that are stimulated by the treatment[6] or by the aggressiveness of the tumor and lack of response to treatment.[7,8] The release of proinflammatory cytokines is typically associated with the onset of fever, which is quickly followed by a cascade of various physical symptoms and feelings of malaise (depressive affect) that typify "sickness behavior".[9] For instance, fever is associated with increased non-rapid eye movement sleep, a period of immune activation both in terms of total time and through slow wave activity.[10]

It is plausible that fever magnifies the relationship between the symptom pair (pain–fatigue/weakness) or the symptom cluster (pain–fatigue/weakness–sleep problems) to feelings of sickness malaise. From a biological perspective, sickness behavior and fever are the end results of a process that is initiated by an endogenous or exogenous pyrogen and that proceeds through a pathway of resultant immune activation, upregulation of lymphocyte transcription factors, and secretion of proinflammatory cytokines. Fever is not only associated with precipitation and perpetuation of proinflammatory cytokines, which occur as immunological reactions to tumors or cancer treatment, but also serves to accelerate internal processes and organ functions.[11] Thus, fever may be a sentinel sign or symptom that operates like a switch, or a sentinel biomarker that is closely tied to the biological process of immune activation and secretion of proinflammatory cytokines that trigger and maintain a cascade of other symptoms. Depressive affect is a useful outcome because it constitutes a holistic marker for sickness malaise, as well as side effects from treatments, analgesics, and smoking cessation.[7]

A parallel, fever-based neurological mechanism is intricately related to the fever-based hematological and immunological mechanism of cytokine production. When macrophages in the bloodstream accelerate cytokine production, the vagus nerve quickly signals the brain to initiate fever through the robust cholinergic anti-inflammatory pathway, which involves numerous nicotinic acetylcholine receptors.[12] Acetylcholine and nicotinic agonists such as nicotine block these receptors and reduce inflammatory cytokines, especially tumor necrosis factor,[13,14] which in turn prevents or attenuates fever,[15,16] neuropathic pain,[17,18] anxiety, and depression.[19] On the other hand, nicotine withdrawal, such as occurs during smoking cessation, aggravates these symptoms.[19]

Fever and other symptoms triggered as side effects of smoking cessation and treatments such as colony-stimulating factor or analgesics may be indistinguishable from those that manifest as sickness behavior resulting from cancer. For instance, in addition to its role in sickness behavior, fever can be produced by opioid medications based on their affinity for opioid mu receptors.[20,21] Similarly, while deterioration in sleep quality, increased daytime fatigue, and increased pain intensity may be mutually reinforcing,[22,23] these effects may stem from uncontrolled pain, and paradoxically, it is speculated, from opioid medications to relieve pain.[24] Higher opioid doses may foster somnolence and sedation, contributing to daytime fatigue and sleepiness.[25–27] In samples of healthy adults and young adults, opioids reduced by 30%–75% the duration of slow wave sleep, the sleep period that promotes feelings of rest and restoration.[28] Opioids also disturb sleep and trigger sleep apnea,[29,30] while disruption of rapid eye movement sleep interferes with drugs demonstrating opioid and serotonin effects.[31]

Biological mechanism(s) that do not necessarily involve fever and occur in the absence of side effects from pain medication or smoking cessation may also influence these symptom pairs or clusters. Many cytokines linked to hypothalamic-pituitary-adrenal axis activation, the autonomic nervous system, and circadian rhythms show explicit diurnal rhythms;[32,33] others alter sleep[34,33] and pain,[35] and are altered by sleep[36,33] and pain.[37] Compared with pain and fever, some fatigue from radiotherapy (such as hormonal or muscular fatigue) does not appear to stem from a steady profusion of cytokines, but undergoes diurnal fluctuation,[38] which may be linked to cytokines with diurnal rhythms, thus dampening the positive linear association between pain and fatigue. This divergence in effect by these different sources of fatigue helps to explain why patients treated with radiation or chemotherapy and attaining relief from pain still experience severe (diurnal) fatigue.[8]

In the absence of specific treatment for fatigue (eg, psychostimulants), the resistance of this subset of fatigue, despite completion of radiation or chemotherapy, may become increasingly pronounced, such that the positive and linear association between pain and fatigue symptoms may be weakened in advanced patients receiving palliative care and disappear altogether during the phase of terminal illness.[39] After diagnosis with breast, colorectal, or prostate cancer, the period prevalence of co-occurring pain, fatigue, and insomnia diminished over the year even as individually occurring symptoms, such as fatigue, may remain pronounced.[40] It is unclear whether this resistant and diurnally fluctuating fatigue should be expected to aggravate, buffer, or have no effect on sensitivity to remaining pain and its relationship to depressive affect, especially if depressive affect stems from depression that interacts with the symptom cluster, triggering or worsening its effect, in contrast with depression that is merely an outcome marker for the motivational state of sickness malaise.

These lines of evidence suggest that in advanced cancer the symptom pair of pain and fatigue, or the symptom cluster of pain, fatigue, and sleep problems, may tap febrile-related processes. It remains unclear whether non-febrile-related processes may be tapped or how. There is an important need to determine whether these paired or clustered symptoms demonstrate similar or different synergistic influences on outcomes such as depressive affect, depending on whether fever is also experienced.

Materials and methods
Sample and measures
The data for the secondary analyses of the current study, collected as part of a primary study funded by the National Cancer Institute (Hospice Program grant, CA48635), involve a sample of 268 individuals with recurrent cancer who were initiating outpatient palliative radiation to reduce bone pain. Medical team providers referred participants from five hospitals in a northeastern US city. Participants were at least 30 years of age, assessed by their oncologists to be beyond cure, although not deemed terminally ill, and had a prognosis of a year or more; they likely differed in diagnosis/treatment stage. Men and women are almost equally represented; their ages range from 30 to 90 years, with half aged 65 years or older. Comorbid health conditions range from none (28.5%), to one (25.8%), to two or more (45.7%). Table 1 reports additional sample characteristics. Participants provided their informed consent, and the University of Pittsburgh internal review board approved the protocol.[41] I have access to a

Table 1 Sample characteristics (n=268)

Characteristic	Frequency	Percentage
Sex		
Female	135	50.3
Male	133	49.7
Age distribution, years		
30–39	7	2.6
40–49	31	11.6
50–59	54	20.1
60–69	94	35.1
70–79	71	26.5
80–90	11	4.1
Primary cancer site		
Breast	58	21.6
Colorectal	13	4.9
Gynecologic	26	9.7
Head and neck	37	13.8
Lung	54	20.2
Prostate	24	9.0
Other	56	20.9
Primary treatment		
Surgery	164	61.2
Curative radiation	83	31.0
Other	21	7.8
Surgery and curative radiation	54	20.1
Comorbid conditions		
Arthritis	73	26.0
Asthma	6	2.1
Diabetes	26	9.3
Emphysema	11	3.9
Heart disease	16	5.7
Hypertension	68	24.2
Arthritis and diabetes	12	4.3
Arthritis and heart disease	6	2.1
Arthritis and hypertension	25	8.9
Arthritis, diabetes, and hypertension	7	2.5
Diabetes and hypertension	12	4.3

Note: Adapted from *Journal of Pain and Symptom Management*, 29(2), Francoeur RB, The relationship of cancer symptom clusters to depressive affect in the initial phase of palliative radiation, 130–155, copyright © 2005, with permission from Elsevier.[7]

version of the initial (baseline) wave of data. The Adelphi University internal review board exempted these data for secondary analysis from review.

Structured interviews with the participants were conducted in their homes, and the same interviewer visited again 4 and 8 months later to conduct repeat interviews. Only 161 participants remained by the third wave of data collection; attrition by 107 participants resulted from death (67%), study withdrawal (18%), being too ill to participate (10.4%), and loss to follow-up (6.6%). The interviewers were trained in correct procedures for administering the structured interview protocol and coding the data.[41] Of particular importance to the current study, this training included emphasis on coding distinguished participant non-response from the response that a symptom did not occur.

The survey[41] included items for participant perception of the degree of difficulty in controlling each of several physical symptoms (each as a single item) during the past month (the Likert-scaled categories are: complete; a lot; some; a little; none). Thus, all symptoms, including the sign of fever, are patient-reported outcomes; objective measures were not also collected. The survey included all 20 items from the Center for Epidemiological Studies-Depression (CES-D) inventory (the four ordinal categories are: rarely; some of the time; much of the time; most of the time). The data afford an opportunity to test whether pain-related interactions with fatigue and sleep problems are also comoderated by fever in predicting depressive affect, a proxy for sickness malaise.

In the current study, the dependent variable of depressive affect during the past week, is an index of five CES-D items of negative affect (ie, sad, felt blue, crying, depressed, lonely), three CES-D items of negative affect within interpersonal and situational contexts (ie, bothered, fearful, thought my life a failure), and three reverse-coded CES-D items of positive affect (ie, hopeful, happy, enjoyed life). CES-D somatic items were excluded because they may constitute symptoms of cancer instead of depression. The internal consistency for the eleven items in these data is very good (α=0.83), and compares favorably with α=0.85 for the entire CES-D.[7] The validity of the depressive affect index is supported by the use of items reflecting positive and negative affect similar to those from two other validated depression scales and by consistent psychometric properties for the constructs of positive and negative affect within the CES-D.[42–52]

Finally, the single-item measures of physical symptoms were initially reported to be common measures derived from previous studies.[41] More recently, a review by Francoeur[7] revealed different lines of converging evidence in the literature that collectively attest to the reliability and validity of self-reported, ordinal, single-item measures of the degree of control across several physical symptoms. All statistical analyses were conducted using Statistical Package for the Social Sciences version 19 software (IBM Corporation, Armonk, NY, USA).

Moderated regression and post hoc assessment of patient profiles

In the current study, moderated regression analyses are conducted to explore associations among responses to symptom items within the clinical sample, based on detecting statistical interactions of mutually influential physical symptoms that synergistically predict depressive affect. A symptom cluster (where all symptoms are endorsed) is subsumed within this broader construct of a symptom interaction effect (which also includes the differential impacts of each comoderating symptom operating alone when the other comoderating symptom does not occur). Also, while symptoms that comprise symptom interactions and clusters occur over the course of the same one-month period in the data, they are based on period prevalence and do not necessarily occur simultaneously. This type of specification allows the moderated regression analyses to incorporate contexts where a prior symptom (eg, fever) that becomes completely controlled may nonetheless trigger other related symptoms or may reveal subgroups with unique symptom management needs (eg, participants with continuing immune activation despite gaining full control over fever). It is not yet clear, however, whether the definition of a symptom cluster (versus interaction) should require all symptoms to occur simultaneously.[53,54]

One accepted definition of a symptom cluster is when concurrent symptoms share a common influence on an outcome.[55] I define symptom pairs and clusters to be limited to physical symptoms (pain, fatigue/weakness, fever, sleep problems) over the past month that are components of the statistical interaction terms, and predict depressive affect. Since the correlative outcome of depressive affect is assessed over the past week, it cannot strictly be part of each symptom cluster (although it becomes a key component of post hoc assessments of patient profiles). In addition, the moderated regression analyses exclude a small subgroup of participants who do not report clinically significant depressive symptoms, based on a total CES-D score between 0 and 10.

These regressions detect, but do not typically foster direct interpretations of, paired or clustered symptom items. Instead, the interaction effect is based on the highest-order interaction term and all derivative terms (ie, all terms representing the lower-order interactions and single variable predictors), allowing us to derive separate patient profiles in subsets of the respondents. These patient profiles may be used to distinguish symptom pairs or clusters from other symptom interaction effects. (A few studies have used patient profiles to dissect the symptom pair of pain and fatigue, based on all subgroup combinations of low and high levels of each symptom;[56,57] in contrast with the current study, these patient profiles do not distinguish the subgroup combinations based on whether they magnify or buffer the relationship between the primary symptom of the cluster and an outcome). In the current study, a post hoc procedure, ie, the extended zero slopes comparison (ZSC), explained and demonstrated in the Supplementary material (Part A), uses the regression slope parameters of each highest-order interaction term, along with

derivative interactions and terms, to foster interpretations of the separate patient profiles. These patient profiles are based on detecting the range of fever values that magnify or buffer the pain–depressive affect relationship at specific values of the other comoderator variable (sleep problems or fatigue/weakness). Also, after removing participants reporting fever, follow-up patient profiles are derived to interpret the nature of comoderation by fatigue/weakness (in place of fever).

The role of fever in explaining the symptom cluster of pain–fatigue/weakness–sleep problems will be examined by including fever initially as a component of the four-way symptom interaction, ie, pain × fever × fatigue/weakness × sleep problems, as well as all derivative lower-order symptom interactions. If the four-way symptom interaction is not statistically significant, or cannot be estimated, models of three-way and derivative interactions involving fever will be estimated instead, and failing that, two-way interactions consisting only of symptom pairs will be estimated. The purpose of this study is to investigate whether fever can be used to distinguish symptom pairs/clusters, and interpret resultant patient profiles, when an original model reveals the four-way interaction, or lower-order derivative interactions involving pain, is statistically significant. Fever that is not fully controlled is hypothesized to accompany fatigue and sleep problems, which may occur as part of physiological acceleration and cytokine deregulation during sickness behavior. However, when fever is not a concern, a review of the literature suggests that fatigue may be more likely to undergo diurnal variation,[38] which may require other approaches, or be more resistant, to palliation;[8,39] the influence of such fatigue in these symptom clusters is apt to differ and will be assessed as well.

Moderated classical regression can be used in small or moderate samples and poses distinct advantages over non-regression methods for assessing population heterogeneity. In the current study, moderated classical regression will be used to detect symptom pairs and clusters. An algorithmic follow-up procedure, ie, the extended ZSC, is applied in the Supplementary material (Part A) to interpret the patient profiles reflected within these symptom pairs and clusters. The regression approach sidesteps the issue of establishing a minimum threshold of symptom expression for determining when a symptom is eligible to contribute to a symptom pair/cluster (see Supplementary material for extended discussions regarding patient profiles [Part A] and the superiority of moderated classical regression over other procedures for symptom cluster research, including the more common approaches of factor analysis, principal components analysis, and cluster analysis [Part B]). This issue is important because even low levels of a symptom could strongly influence a symptom cluster, for instance, participants reporting a lot (but not complete) control over fever are still implied to have an activated inflammatory pathway.

In the current study, original symptom values are initially rescaled to be mean-centered to allow for more meaningful interpretation of the findings from the moderated regressions (ie, when remaining symptoms are at their mean values rather than at zero). Each raw regression is re-estimated using sequential residual centering (SRC), a sequential application of residual centering developed and validated by the author.[3,58] This innovation is an extension to residual centering for reducing multicollinearity, described by Lance,[59] where each mean-centered variable also becomes residual-centered.

Results

Frequencies of depressive affect and physical symptoms are reported in Tables 2 and 3. All symptom distributions are

Table 2 Extent of symptom control (n=268)

Symptom	Mean [mode] (SD)[a]	Does not occur n (%)	Complete (0) n (%)	A lot (1) n (%)	Some (2) n (%)	Little (3) n (%)	None (4) n (%)
Change in bowel habits	0.94 [1] (1.40)	145 (54.1)	13 (4.9)	48 (17.9)	19 (7.1)	8 (3.0)	35 (13.1)
Fatigue/weakness	1.62 [1] (1.49)	67 (25.0)	10 (3.7)	79 (29.5)	35 (13.1)	23 (8.6)	54 (20.1)
Fever	0.25 [1] (0.87)	238 (88.9)	3 (1.1)	12 (4.5)	1 (0.4)	3 (1.1)	11 (4.1)
Nausea/vomiting	0.83 [4] (1.41)	175 (65.3)	4 (1.5)	34 (12.7)	14 (5.2)	5 (1.9)	36 (13.4)
Pain	1.19 [1] (1.45)	120 (44.8)	6 (2.2)	55 (20.5)	36 (13.4)	10 (3.7)	41 (15.3)
Poor appetite	1.25 [4] (1.58)	140 (52.2)	8 (3.0)	19 (7.1)	36 (13.4)	18 (6.7)	47 (17.5)
Shortness of breath/ difficulty breathing	0.68 [1] (1.29)	188 (70.1)	3 (1.1)	33 (12.3)	12 (4.5)	4 (1.5)	28 (10.4)
Sleep problems	1.25 [4] (1.66)	148 (55.2)	7 (2.6)	23 (8.6)	16 (6.0)	17 (6.3)	57 (21.3)
Weight loss	1.21 [4] (1.64)	144 (53.7)	13 (4.9)	25 (9.3)	15 (5.6)	17 (6.3)	54 (20.1)

Notes: [a]For the purpose of estimating symptom means, modes, and standard deviations, symptoms that do not occur are coded into the category for "complete control" (=0). Adapted from *Journal of Pain and Symptom Management*, 29(2), Francoeur RB, The relationship of cancer symptom clusters to depressive affect in the initial phase of palliative radiation, 130-155, copyright © 2005, with permission from Elsevier.[7]
Abbreviation: SD, standard deviation.

Table 3 Extent of depressive affect and frequencies of symptom interactions (n=268)

Depressive affect	11[a] n (%)	12–14 n (%)	15–17 n (%)	18–38 n (%)
Mean: 15.54 Actual range: 11–38 Possible range: 11–44	68 (25.5)	79 (29.6)	50 (18.7)	70 (26.2)
Symptom interactions	**Frequency when incomplete control (a lot =1 to none =4) of each component n (%)**			
Pain × fatigue/weakness	110 (41.0)			
Pain × sleep problems	141 (52.8)			
Fatigue/weakness × sleep problems	190 (71.2)			
Pain × fatigue/weakness × sleep problems	110 (41.2)			

Notes: [a]Depressive affect is an index of five CES-D items of negative affect (ie, sad, felt blue, crying, depressed, lonely), three CES-D items of negative affect within interpersonal and situational contexts (ie, bothered, fearful, though my life a failure,), and three reverse-coded CES-D items of positive affect (ie, hopeful, happy, enjoyed life). The lowest possible score is eleven, resulting when all eleven CES-D items are endorsed as "rarely" occurring whereas the highest possible score is 44, resulting when all eleven CES-D items are endorsed as "most of the time". Scores are reported in ranges representing similar numbers (n) of participants (ie, 50 to 79) that make up similar percentages of the total sample (ie, 18.7% to 29.6%). Adapted from *Journal of Pain and Symptom Management*, 29(2), Francoeur RB, The relationship of cancer symptom clusters to depressive affect in the initial phase of palliative radiation, 130-155, copyright © 2005, with permission from Elsevier.[7]
Abbreviation: CES-D, Center for Epidemiologic Studies-Depression.

highly skewed, with most participants reporting complete control of each symptom.

These linear effects of common symptoms, quadratic (curvilinear) effects of specific symptoms that are components of symptom interactions, and specific symptom interactions together predict depressive affect in regressions reported in Table 4. As expected, relevant parameter estimates are identical in the raw and SRC regressions; inflated variance inflation factor (VIF) values in the raw regression fall dramatically to essential VIF (EVIF) values less than 10 in the SRC regressions. (In Table 4, cell entries appear in bold when VIF values fall dramatically after SRC and the b parameter becomes newly statistically significant.) Thus, none of the predictors in the SRC runs are identified to be associated with problematic multicollinearity.[60] In addition to meeting the common standard that all VIFs (here, EVIF values) be less than 10, the EVIF in regressions 1A, 1B, 2, and 3 all meet the more conservative rule that the mean of all VIFs (here, EVIF values) from each regression must not be considerably larger than 1;[61] however, the mean value of 2.6 in the SRC run for regression 4 suggests that while multicollinearity is dramatically reduced, the remaining multicollinearity due to essential ill conditioning could still have limited influence.

Table 4 SRC-moderated regressions were re-estimated twice, replacing mean centering with mode centering and a second alternative for centering based on the ordinal category with the second-highest frequency of responses. In the Supplementary material (Part A), I report only the post hoc analyses based on the original mean centered estimates; however, findings remain very similar across all three centering options.

The Supplementary material (Part C) discusses critical advantages of SRC over raw regression in the context of mean-centered or other-centered variables. Collectively, SRC results in new (or more highly) significant effects, based on essential standard error (ESE) parameters, than the raw regression in Table 4 (even as the relevant b and SE parameters remain unchanged). In SRC runs, pain × fatigue/weakness × sleep and pain × sleep, significant at $P<0.05$ or $P<0.01$ in descriptive and explanatory models (regressions 1A and 1B) that test three-way (second-order) interactions separately, become non-significant when all interactions are tested simultaneously (regression 4). However, pain × fever, pain × fever × sleep, and pain × fever × fatigue/weakness, significant at $P<0.05$ or $P<0.01$ in separate three-way (second-order) explanatory models (regressions 2 and 3), all become highly significant ($P<0.001$) in this simultaneous model (regression 4).

Post hoc analyses in the Supplementary material (Part A) reveal fever magnifies the pain–depressive affect relationship when there is little or no control over sleep problems or less than full control over fatigue/weakness (ie, a lot of control, a little control, no control). Furthermore, when fever presents, specific ranges of the other co-occurring symptoms (sleep problems or fatigue/weakness) also magnify the pain–depressive affect relationship. Considering both magnifier effects together, there exists a mutually synergistic and compounded magnifier effect on the pain–depressive affect relationship in the context when fever presents within specific ranges of either sleep problems (a little or no control) or fatigue/weakness (a lot of control, a little control, no control). The relationship is buffered in the lower ranges of these two symptoms where they are better controlled.

Nausea and breathing difficulties are added to these explanatory models because in previous secondary analyses with these data, these common symptoms were revealed to be components of symptom interactions also involving pain or fatigue/weakness, which could overlap those in the current study.[7] It should also be noted that interpretations of the findings did not change, or changed in minor ways, when each explanatory regression also included statistical control

Table 4 Depressive affect predicted by physical symptoms and symptom interactions[a]

Independent variables	Unstandardized b (SE) (ESE: SE from essential ill-conditioning only; reported if VIF >10) [VIF: VIF from essential- and inessential-ill conditioning; reported if VIF >10] [EVIF: VIF from essential ill-conditioning only; reported if VIF >10]				
	1A[b] Pain × fatigue/weakness × sleep problems: descriptive model	1B[b] Pain × fatigue/weakness × sleep problems: explanatory model	2 Pain × sleep problems × fever: explanatory model	3[c] Pain × fever × fatigue/weakness: explanatory model	4[d] All four 3-way interactions specified simultaneously: explanatory model
Pain	0.267 (0.328)	0.240 (0.332)	0.474 (0.360)	0.125 (0.359)	-0.295 (0.478)
Shortness of breath, difficulty breathing		-0.068 (0.233)	-0.130 (0.245)	-0.200 (0.245)	-0.287 (0.248)
Sleep problems	-0.093 (0.344)	-0.104 (0.343)	0.102 (0.360)	0.506 (0.195)***	0.558 (0.463)
Nausea, vomiting		0.595 (0.215)**	0.735 (0.231)**	0.732 (0.225)***	0.831 (0.234)****
Fever		0.022 (0.343)	0.414 (0.343)	0.203 (1.430) (ESE: 0.343) [VIF: 18.235] [EVIF: 1.051]	-0.746 (1.623) (ESE: 0.445)@ [VIF: 15.891] [EVIF: 1.195]
Fatigue/weakness	0.614 (0.249)*	0.397 (0.261)	0.270 (0.232)	0.212 (0.267)	0.123 (0.320)
Pain²	0.110 (0.193)	0.103 (0.192)	0.192 (0.196)	0.271 (0.194)	0.186 (0.203)
Sleep problems²	0.400 (0.221)@	0.389 (0.220)@	0.421 (0.226)		0.524 (0.238)*
Fever²			-0.344 (0.506) (ESE: 0.125)** [VIF: 24.160] [EVIF: 1.468]	0.010 (0.484) (ESE: 0.119) [VIF: 22.506] [EVIF: 1.367]	0.587 (0.731) (ESE: 0.210)** [VIF: 33.132] [EVIF: 2.741]
Fatigue/weakness²	-0.081 (0.141)	0.078 (0.178)		0.115 (0.182)	0.052 (0.187)
Pain × sleep problems	0.013 (0.177)	-0.391 (0.147)**	-0.228 (0.128)		-0.073 (0.162)
Pain × fever			0.445 (0.422)	-1.190 (0.477)*	-3.448 (1.404)* (ESE: 0.377)**** [VIF: 58.853] [EVIF: 4.242]
Pain × fatigue/weakness	-0.402 (0.147)**	-0.093 (0.140)		-0.220 (0.137)	0.193 (0.239)
Fever × sleep problems			0.455 (0.354)		2.919 (1.308)* (ESE: 0.388)**** [VIF: 53.274] [EVIF: 4.684]
Sleep problems × fatigue	-0.047 (0.130)	-0.024 (0.130)			-0.508 (0.234)*

(Continued)

Table 4 (Continued)

Independent variables	Unstandardized b (SE) (ESE: SE from essential ill-conditioning only; reported if VIF >10) [VIF from essential- and inessential-ill conditioning; reported if VIF >10] [EVIF: VIF from essential ill-conditioning only; reported if VIF >10]				
	IA[b] Pain × fatigue/weakness × sleep problems: descriptive model	IB[b] Pain × fatigue/weakness × sleep problems: explanatory model	2[b] Pain × sleep problems × fever: explanatory model	3[c] Pain × fever × fatigue/weakness: explanatory model	4[d] All four 3-way interactions specified simultaneously: explanatory model
Fever × fatigue/weakness				−0.325 (0.391)	−0.658 (0.749) (ESE: 0.302)* [VIF: 10.336] [EVIF: 1.676]
Pain × sleep problems × fever[f]			−0.361 (0.173)*		−0.325 (0.327) (ESE: 0.077)**** [VIF: 20.574] [EVIF: 1.153]
Pain × fatigue/weakness × sleep problems	0.128 (0.078)+ (ESE: 0.054)* [VIF: 2.132] [EVIF: 1.009]	0.118 (0.078)++ (ESE: 0.053)* [VIF: 2.152] [EVIF: 1.019]			−0.001 (0.076)
Pain × fever × fatigue/weakness[e]				0.660 (0.243)**	2.105 (0.801)*** (ESE: 0.294)**** [VIF: 60.049] [EVIF: 8.067]
Fever × fatigue/weakness × sleep problems					−1.528 (0.837)@ (ESE: 0.277)**** [VIF: 71.800] [EVIF: 7.819]
R^2, F value	0.164, 4.978****	0.190, 4.518****	0.194, 4.663****	0.210, 5.153****	0.237, 3.71 |****

Notes: n=268; @P<0.10; *P<0.05; **P<0.01; ***P<0.005; ****P<0.001 (all tests are two-tailed); +P=0.102. Due to this tentative level of statistical significance in the raw regression for 1A, VIF, essential VIF, and essential SE are reported despite that VIF was <0; P<0.05 after SRC. ++P=0.128. Due to this tentative level of statistical significance in the raw regression for 1B, VIF, essential VIF, and essential SE are reported despite that VIF was <10; P<0.05 after SRC. Cell entries in bold show dramatic reductions in inessential multicollinearity (compare VIF and essential VIF) and statistically significant b parameters. [a]As a general rule according to Belsley et al,[60] the VIF should not exceed 10. Entries for a predictor are in bold when statistically non-significant b parameters in the raw regression (using SE) become significant in the SRC run (ie, using essential SE) at P<0.05 or below, or when significant b parameters in the raw regression meet the threshold for statistical significance at a lower P-value in the SRC run; [b]explanatory regressions are reported in 2 through 4. Statistically significant interactions in 2 and 3 remain significant in corresponding *descriptive* regressions where the only predictors are the interactions, their derivative main-effect components, and related curvilinear terms. Thus, just as descriptive regression 1A provides evidence that the pain × fatigue/weakness × sleep problems interaction can be detected within the data, and explanatory regression 1B reveals that this interaction cannot be attributed to other symptoms (ie, it remains statistically significant), similar inferences can be made for the interactions tested in 2 and 3; [c]separate regressions to test fever × fatigue/weakness × sleep problems and pain × fever × fatigue/weakness × sleep problems (not shown) did not reveal these interactions to be statistically significant. The coefficient of the four-way interaction switches sign (from positive to negative) and becomes significant only after excluding 13 influential outliers; the moderate sample size may contribute to its lack of significance in the full sample. Thus, only up to three-way (second-order) regression model specifications can be taken to be valid for use with these data; [d]influential observations with Cook's d values greater than 4/n, or 0.140, were dropped. Two observations were dropped in 1, one dropped in 2 and in 3, and seven dropped in 4. In 4, sleep problems², pain × fever, pain × sleep problems × fever and pain × fever × fatigue/weakness remain significant in the SRC regression when all seven outliers are retained; [e]in 4, the regression specification before the last interaction is added (ie, fever × fatigue/weakness × sleep problems), the parameters for pain × fever × fatigue/weakness are statistically significant (b=0.750, essential SE=0.111****, essential VIF=1.146). The inclusion of this last interaction term in the regression serves to dramatically increase the b parameter value for pain × fever × fatigue/weakness (b=2.105, essential SE=0.294****, essential VIF=8.067). Thus, pain × fever × fatigue/weakness is based, in part, on a "suppression effect". When 4 is rerun with all observations (ie, including the influential cases), the suppression effect remains as well [ie, compare the runs: 1) with fever × fatigue/weakness × sleep problems: b=7.250, essential SE=7.562 (non-significant), essential VIF=60.049; and 2) without fever × fatigue/weakness × sleep problems: b=2.582, essential SE=1.128*, essential VIF=8.529]. The highly inflated essential VIF value of 60.049 in (1) happens to be identical to the VIF value for the same term in the raw regression that includes only non-influential observations (ie, see regression 4). Thus, adding the influential observations simply adds back the inessential multicollinearity removed by SRC; however, this multicollinearity now occurs within the same observations (not just between the two interaction terms) and thus is now essential multicollinearity; [f]in 4, the inclusion of fever × fatigue/weakness × sleep problems creates a less dramatic suppressor effect on pain × fever × fatigue/weakness × sleep problems: b=−0.325, essential SE=0.077****, essential VIF=1.153; and 2) without fever × fatigue/weakness × sleep problems: b=−0.199, essential SE=0.076***, essential VIF=1.101. When outliers are excluded, we can compare the runs: 3) with fever × fatigue/weakness × sleep problems: b=−0.325, essential SE=−0.137**, essential VIF=1.153; and 4) without fever × fatigue/weakness × sleep problems: b=−0.199, essential SE=−0.167 (non-significant), essential VIF=1.101. Comparing (1) with (3) and (2) with (4), the respective b parameters and essential VIF values do not change; however, the essential SE values do change, such that pain × fever × sleep problems ends up becoming non-significant when outliers are included. This lack of statistical significance occurs in the context of the highly inflated essential VIF value of 60.049 for pain × fever × fatigue/weakness, which also became non-significant (refer to footnote e). These findings reflect overlapping variation across all three third-order interactions that is contributed by the set of influential observations and constitutes essential multicollinearity. Copyright © 2014. Adapted from Dove Medical Press. Francoeur RB. Using an innovative multiple regression procedure in a cancer population (Part 1): detecting and probing relationships of common interacting symptoms (pain, fatigue/weakness, sleep problems) as a strategy to discover influential symptom pairs and clusters. *Onco Targets Ther*. 2015;8:45–56.[3] Adapted from Francoeur RB. Could sequential residual centering resolve low sensitivity in moderated regression? Simulations and cancer symptom clusters. *Open Journal of Statistics*. 2013;3:24–44.[58]

Abbreviations: SE, standard error; ESE, essential SE from essential ill-conditioning only; VIF, variance inflation factor; EVIF, essential variance inflation factor from essential ill-conditioning only; SRC, sequential residual centering.

variables for sex (a dummy variable representing males), age (<65 years versus 65+ years), an ordinal variable for illness comorbidity (none, one, two or more conditions), and a series of dummy variables selecting out participants who did not experience any given symptom (see Supplementary material [Part C]). The dummy variables were added to prevent conflation, especially between absence of fever and control of fever. (Since fever is an end result of an inflammatory pathway, controlling the fever that results still would imply that the inflammatory pathway has been activated. However, only three participants reported experiencing fever in the past that was completely controlled during the past week, which explains why findings are similar when either the absence or control of fever is specified.)

I now review each regression separately. The SRC descriptive (1A) and explanatory (1B) regressions result in statistical significance of pain × fatigue/weakness × sleep at $P<0.05$ in Table 4.

Regressions 2 and 3 in Table 4 are explanatory analyses when fever is tested as part of a statistically significant symptom interaction predicting depressive affect. The unreported four-way (third-order) interaction, pain × fever × fatigue/weakness × sleep, is not statistically significant; however, separate regressions reveal two significant three-way interactions: pain × sleep × fever in regression 2; and pain × fever × fatigue/weakness in regression 3. The nature of the symptom interaction effects in regressions 2 and 3 are probed in post hoc analyses described in Supplementary material (Part A). Interpretations are reported there and in Table 5.

The exhaustively specified explanatory regression 4 in Table 4 comprises the fully specified model of all four three-way (ie, second-order) interactions across the four symptoms of pain, fever, fatigue/weakness, and sleep. Three of the interactions (pain × sleep × fever; pain × fever × fatigue/weakness; and fever × fatigue/weakness × sleep problems) become highly significant ($P<0.001$) in the SRC regression.

Given the prominence of fever in regressions 2–4, the original descriptive (1A) and explanatory (1B) regressions are re-estimated separately within the subgroup in which fever does not occur (238 participants). Results are similar for 1A and 1B and so I report findings here only for 1B (for which

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Table 5 Interpretations of comoderator effects detected in reported regressions

Regression 2: testing pain × sleep problems × fever

A. Comoderation by fever

When there is a little control or no control over sleep problems (w=3 or 4), fever magnifies the pain–depressive affect relationship over the full range of fever, from complete control to no control (ie, z=0 to 4; n=74).

B. Comoderation by sleep problems

1. When there is some control to no control of fever (z=2, 3, or 4), sleep problems magnify the pain–depressive affect relationship over the full range of sleep problems, from complete control to no control (ie, w=0 to 4; n=15).

2. When there is complete control of fever (z=0), sleep problems magnify the pain–depressive affect relationship over the range of sleep problems from complete to some control (ie, w=0 to 2; n=182) and buffer the pain–depressive affect relationship over the range of sleep problems from a little control to no control (ie, w=3 and 4; n=58).

   • When there is a lot of control of fever (z=1), sleep problems magnify the pain–depressive affect relationship at complete control of sleep problems (ie, w=0; n=3) and buffer the pain–depressive affect relationship over the range of sleep problems from a lot of control to no control (ie, w=1 to 4; n=9).

Regression 3: testing pain × fever × fatigue/weakness

A. Comoderation by fever

1. When there is a little control or no control of fatigue/weakness (w=3 or 4), fever magnifies the pain–depressive affect relationship over the range of fever, from a lot of control to no control (ie, z=1, 2, 3, 4; n=14).

   • When there is a lot of control of fatigue/weakness (w=1), fever magnifies the pain–depressive affect relationship over the full range of fever, from complete to no control (ie, z=0 to 4; n=7).

B. Comoderation by fatigue/weakness

2. At both extremes, when there is either complete control or no control of fever (z=0 or 4), fatigue/weakness buffers the pain–depressive affect relationship at complete control to a lot of control of fatigue/weakness (w_{low} =0 and 1; n=149) and magnifies the pain–depressive affect relationship at some control to no control of fatigue/weakness (w=2, 3, 4; n =102).

   • When there is a lot of control to a little control of fever (z=1, 2, or 3), fatigue/weakness buffers the pain–depressive affect relationship at complete control to some control of fatigue/weakness (w_{low} =0, 1, and 2; n=9) and magnifies the pain–depressive affect relationship at a little control to no control of fatigue/weakness (w=3, 4; n=7).

Regression 1B: follow-up: pain × fatigue/weakness when fever control is not concern

A. When fever control is not a concern, fatigue/weakness magnifies the pain–depressive affect relationship at complete to a lot of control of fatigue/weakness (ie, w=0, 1; n=146) and buffers the pain–depressive affect relationship at some to no control of fatigue/weakness (ie, w=2, 3, 4; n=94).

Note: Adapted from Francoeur RB. Could sequential residual centering resolve low sensitivity in moderated regression? Simulations and cancer symptom clusters. *Open Journal of Statistics.* 2013;3:24–44.[58]

R^2=0.222 and F=4.541; a follow-up interpretation reported in Table 5 is derived in Supplementary material (Part A) using the extended ZSC). In the subgroup in which fever does not occur, the slope for pain × fatigue/weakness × sleep problems deteriorates, switches sign, and becomes statistically insignificant (1B: b=−0.004, SE =0.077, z=−0.052, P=0.958, VIF =1.770), while the slope for the derivative interaction, pain × fatigue/weakness, increases and becomes newly significant (1B: b=−0.328, SE =0.142, z=−2.305, P<0.05, VIF =1.334). In the upper range, fatigue/weakness now switches from magnifying to buffering the pain–depressive affect relationship. The slope of the quadratic (curvilinear) term for sleep, sleep problems^2, also becomes newly significant (1B: b=0.499, SE =0.245, z=2.039, P<0.05, VIF =4.410).

Finally, correlations among pain, fatigue/weakness, and sleep are somewhat stronger in those with incomplete fever control than the remaining sample without fever or complete fever control (pain and fatigue, 0.444 versus 0.291; pain and sleep, 0.488 versus 0.337; fatigue/weakness and sleep, 0.427 versus 0.333). Similar associations occur when the three participants with completely controlled fever are classified with those experiencing incomplete control.

Discussion
Clustering responses to symptom items: detection using moderated regression

The statistical significance of the three-way interaction (pain × fatigue/weakness × sleep problems) in the entire sample, but not in the subgroup without fever, along with follow-up regressions 2–4 in Table 4, suggest that fever serves as a sentinel sign of a unique and possibly distinct subgroup in which these three interacting physical symptoms occur in the context of sickness malaise. This subgroup may consist of participants experiencing neutropenic fever, infection, or a single disease type, and/or receiving a concurrent medication that induces fever, such as a colony-stimulating factor, or perhaps for a greater number of cases, the use of opioids to relieve pain. Limitations of the data do not permit investigation of these possibilities (see Supplementary material, Part D). However, participants experiencing fever appear to represent a homogeneous subgroup, compared with the larger subsample not experiencing fever, since they reveal stronger associations among pain, fatigue/weakness, and sleep problems. (These findings, of course, should not be misconstrued as ruling out possibilities for statistical significance of the three-way interaction within more finely distinguished subgroupings of outpatients without fever.)

In Table 4, the unstandardized b values in regressions 1A and 1B for pain × fatigue/weakness × sleep problems are smaller in magnitude than the remaining fever-based interactions in regressions 2 and 3. Furthermore, in the exhaustive and simultaneous model (regression 4), two of the terms (pain × sleep problems and pain × fatigue/weakness × sleep problems) become statistically insignificant, whereas the three three-way interactions involving fever become highly significant (P<0.001). This pattern of findings suggests there may be some "noise" in the relationship between pain × sleep problems × fatigue/weakness and depressive affect, which could be due to diurnal fluctuation in sources of fatigue that are unrelated to fever, and by extension, that are unrelated to system-wide physiological acceleration, proinflammatory cytokines, and sickness behavior.[38] This pattern of findings is also consonant with the perspective that the symptom cluster of pain, fatigue, sleep, and depressive affect may be less pronounced in more progressed phases of illness.[39]

To test this possibility, I conducted follow-up runs of regressions 1A and 1B, selecting out the participants reporting fever. In the remaining sample of 238 participants, pain × sleep problems × fatigue/weakness is no longer statistically significant while the derivative interaction of pain × fatigue/weakness, and the quadratic (curvilinear) term for sleep problems, become newly significant. Thus, when especially pronounced, sleep problems independently predicts depressive affect, and not necessarily in the same participants for whom concurrent pain and fatigue/weakness uniquely predict depressive affect.

Since sleep problems is no longer an interacting component of the symptom cluster, it may be a salient symptom only in participants with uncontrolled fever who experience sickness malaise (indicated by depressive affect). This lack of involvement of sleep problems in participants without fever could mean that the significant interaction of pain × fatigue/weakness in these participants is not characterized by the same unifying mechanism of system-wide physiological acceleration and steady profusion of proinflammatory cytokines associated with symptom precipitation and sickness malaise, a continual process expected to lead to, or perhaps be worsened by, insomnia or sleep disruption. However, it is important to stress that the outcome measure of depressive affect over the past week cannot be assured to reveal only sickness malaise. In some cases, it may reflect pre-existing dysthymia or other depression, although these participants might still be aware of more frequent or worsening depression symptoms during the past month, which may stem from their recent experiences of pain and physical symptoms.

Also, depressed participants may be less likely to report the fever characteristic of sickness malaise, as suggested in the next section on patient profiles (see Participants without fever section). In any event, findings should be interpreted with caution.

I now turn back to Table 4 to compare the remaining original regressions. When tested alone or with the remaining three-way interactions and their components, the robust statistical significance of pain × sleep problems × fever (regressions 2 and 4); and pain × fever as well as pain × fever × fatigue/weakness (regressions 3 and 4) suggests that all four symptoms (pain, fever, fatigue/weakness, sleep problems) may operate as triggers of each other, a notion supported further by the two additional newly significant second-order interactions of sleep problems × fever and sleep problems × fatigue in the exhaustively specified model (regression 4). Stated differently, there is no evidence that any of these symptoms occur merely as reactions or consequences of other symptoms without also influencing reactive feedback effects. The robustness of fever in three of the four statistically significant three-way symptom interactions in regression 4 raises the issue about the extent to which prior findings in the literature that support pain, fatigue, and sleep disturbance, but do not also consider fever, may actually be capturing the unmeasured impact of uncontrolled fever, which the theory of sickness behavior identifies as part of the primary trigger, or at least strongly associated with the primary trigger, of pain, insomnia or sleep disruption, and fatigue/weakness. For instance, the bidirectional relationship between pain and sleep disturbance[62–64] may be a reactive feedback effect that is triggered or maintained by fever or associated cytokine processes, which is implied in the current study by the interaction pain × sleep problems × fever and its derivative interactions.

Patient profiles: interpretation using a post hoc procedure

Participants with fever

The post hoc analyses of patient profiles provide further insights into these symptom interactions. Interpretation 1 for regression 2 in Table 3 reveals that diminishing control of fever across its full range (complete to no control) magnifies the pain–depressive affect relationship when there is only a little control, or no control, over sleep problems. This outcome suggests that incomplete control of fever may be: a trigger or represent a chronic component of sleep problems that is associated with sickness malaise (ie, a direct mechanism); and/or highly associated with problematic component(s) like

pain, fatigue, and weakness that also influence (and may be influenced by) sleep problems and sickness malaise (ie, an indirect mechanism). Partial support for the direct mechanism can be inferred from an experimental study; sleep restriction over a 12-day period increased the proinflammatory cytokine interleukin-6, which in turn precipitated or magnified pain.[62] Several inflammatory markers became elevated; however, fever was not tested as one of them, perhaps because the study involved healthy volunteers. Nevertheless, the role of fever during sickness in precipitating the release of proinflammatory cytokines such as interleukin-6 suggests it could serve as a trigger (or be closely tied to a biological trigger) or a chronic component of restricted sleep.

Interpretation 1 for regression 3 in Table 5 indicates that over most of its range (a lot of control to no control), fever magnifies the pain–depressive affect relationship when there is only a little control or no control over fatigue/weakness. Interpretation 2 for regression 3 in Table 5 reveals a similar magnifier effect occurs across the full range of fever (complete to no control) when there is a lot of control of fatigue/weakness. These outcomes appear to support the classic mammalian response of sickness behavior in which proinflammatory cytokines trigger a cascade of symptoms that immobilize the organism, thus conserving and redirecting energy into a strong immune response. They may also reflect analgesic side effects such as opioid-related fever and sedation.

The patient profiles for pain × sleep problems × fever and for pain × fever × fatigue/weakness reveal that fever worsens the pain–depressive affect relationship when there is no control either of sleep problems or fatigue/weakness. Curiously, in Table 4, fever does not appear to magnify the relationship of fatigue-depressive affect or sleep problems–depressive affect (ie, fever × fatigue/weakness and fever × sleep problems are not statistically significant) without also involving pain. Therefore, considering all significant two-way and three-way interactions, it may be speculated that the findings suggest the magnifier effect of fever on the pain–depressive affect relationship could be occurring through the impact of fever in increasing pain sensitivity, which in turn may contribute to fatigue/weakness and sleep problems; and in increasing sleep problems, which in turn may contribute to pain sensitivity.

This context suggests that crossover effects of fever interventions may relieve these remaining symptoms as well. As explained earlier in the review of the literature, the vagus nerve and its cholinergic anti-inflammatory pathway appears to be a useful, widely networked route to relieve fever,[12–16] neuropathic pain,[17,18] anxiety, and depression.[19] Moreover, relieving fever may slow the rate of catabolism in

autoimmune diseases such as cancer,[65] perhaps even slowing the progression to catabolic wasting (cachexia). Complementary approaches that stimulate the vagus nerve, acetylcholine, and the cholinergic anti-inflammatory pathway show promise in relieving these symptoms, and include exercise, electroacupuncture, hypnosis, meditation, behavioral conditioning, and biofeedback.[14,66–68] Pharmaceutical companies are developing agents based on promising non-addictive nicotinic agonists to address autonomic dysfunction of the cholinergic anti-inflammatory pathway.[14]

As expected, nicotine from tobacco also stimulates the cholinergic anti-inflammatory pathway and reduces fever,[69] while nicotine withdrawal leads to fever, anxiety, and depression.[19] Thus, fever-related symptom clusters involving pain and predicting depression may especially of concern in patients who have reduced or quit smoking, perhaps in response to advanced disease. Periodic follow-up appears warranted to reassess symptoms and to encourage efforts and alternatives for smoking cessation (eg, use of a nicotine inhaler, patch, spray, or gum). Patient education and motivational interviewing for smoking cessation should address the role of complementary medical procedures, not only in relieving nicotine withdrawal, but in preventing and counteracting precipitation of cancer symptoms that may occur simultaneously and afterwards. In the future, new drugs based on non-addictive nicotinic agents may also become options.

Participants without fever

Next I interpret the patient profile reflected by the post hoc analysis of follow-up regression 1B within the subsample of participants without fever. (Recall that pain × fatigue/weakness, but not pain × fatigue/weakness × sleep problems, is statistically significant within this subgroup in follow-up runs of regressions 1A and 1B). The pain–depressive affect relationship is magnified along the range of complete to a lot of control over fatigue/weakness but then is buffered along the upper range of fatigue/weakness control problems (ie, some to no control). This checkered patient profile implies that pain and fatigue/weakness in these participants might not stem from the same underlying mechanism. This non-constant relationship also raises the possibility that although pain and fatigue/weakness occur during the same one-week period, they do not necessarily occur simultaneously; either or both of these symptoms could fluctuate diurnally, as Hickok et al[8] highlighted.

Moreover, while occasional fatigue and weakness may reflect dynamics that worsen the relationship between pain

and depressive affect, higher levels of fatigue and weakness may either represent a threshold "ceiling" effect or reveal adaptation within this relationship, perhaps by interfering with awareness of painful sensations or the experience of depressive affect. The range of symptoms in which symptom clusters reveal buffering effects suggests there may be potential for concurrent symptoms to protect against higher levels of either physical symptoms or depressive affect, or alternatively, concurrent symptoms may simply interfere with symptom reporting. In these situations, symptom-specific interventions with crossover effects could end up removing this protection and worsening remaining symptoms in the cluster. Yet another explanation for the switch to buffering effects when fatigue and weakness are less controlled could be the impact of resistant fatigue, as Kaasa et al[39] revealed, and to which patients may gradually adapt. This last scenario would suggest a tradeoff between the degrees of success in controlling pain versus controlling pronounced fatigue/weakness. Conversely, the magnifier effect along the lower range of fatigue/weakness (ie, complete to a lot of control) suggests that there may still be potential for individual interventions to relieve pain and occasional, non-resistant fatigue.

The buffering effect should not be dismissed prematurely, however, as failing to suggest a promising context for crossover intervention effects. On the contrary, it may serve as a clue that the potential context for desirable crossover effects could require that the intervention operate more broadly by also directly relieving depressive affect (in addition to its indirect impact on depressive affect from relieving the physical symptoms). A body of empirical evidence distinguishes the malaise in sickness behavior, a motivational state that manifests as the body shifts its priorities to cope with inflammation, from the clinically significant depression that occurs not only as an outcome of physical symptoms but precipitates and exacerbates them as well (see overview by Dantzer and Kelley[70]). Similarly, in the current study, the portion of depressive affect that would need to be directly relieved by a crossover intervention stems from a bidirectional relationship where depressive affect, and, by implication, clinically significant depression, is not only an outcome of physical symptoms, but also precipitates them. The synergistic effects of the bidirectional relationship of depressive affect on and by physical symptoms may be modeled by an interaction term in which depressive affect and physical symptoms are components.

We can test the merit of such a clue by incorporating depressive affect as an additional symptom within the physical symptom cluster, and finding that the resulting cluster

reveals magnifier effects in predicting another important, and more distal, outcome related to functional status. Indeed, in the same secondary data from non-febrile participants, but not in the overall sample (which includes those with fever), I subsequently detected two symptom clusters: pain–fatigue/weakness–depressive affect (b=0.005, ESE =0.0015; z=2.751; P<0.01); and pain–fatigue/weakness–sleep problems–depressive affect (b=0.004; ESE =0.0008; z=4.942; P<0.001), in which depressive affect is one of the cluster components that predicts a single-item measure of mobility problems. In both clusters, each of the comoderating symptoms magnified the pain–mobility problems relationship.

This wider context of comagnifying effects, where depressive affect is incorporated as an additional symptom within the physical symptom cluster, suggests that crossover interventions for each symptom cluster could yield direct relief of depressive affect, and depression, in non-febrile patients. In febrile patients, on the other hand, there is no evidence for a bidirectional relationship between these symptom clusters and depressive affect since both clusters are no longer statistically significant once participants with fever are added. This finding can also be taken as providing further indirect support for the unidirectional relationships of the earlier fever-based symptom clusters to sickness malaise (in contrast to bidirectional relationships with depressive affect from depression).

These findings afford an opportunity to integrate related evidence from the earlier literature review. Recall that in the absence of fever, fatigue may be more likely to undergo diurnal variation,[38] which may require other approaches, or be more resistant, to palliation.[8,39] This finding suggests an explanation why mobility problems are predicted by the two symptom clusters (pain–fatigue/weakness-depressive affect and pain–fatigue/weakness–sleep problems–depressive affect): diurnal variation in fatigue and weakness may lead directly to mobility problems by making it difficult to stand and walk reliably or long enough to complete activities of daily living, which may be worsened (ie, comagnified) by the remaining clustering symptoms (ie, by physical limitations from pain, lack of motivation suggested by depressive affect, and tiredness from insufficient sleep). It is unclear, however, whether fatigue/weakness or pain should be considered the primary symptom that limits mobility, which may differ from person to person. For some participants, pain may be the primary symptom that limits mobility, with fatigue/weakness, as it waxes diurnally, operating instead to comagnify the impact of pain.

Let us return to the discussion of comagnifying effects within a symptom cluster as a context in which crossover
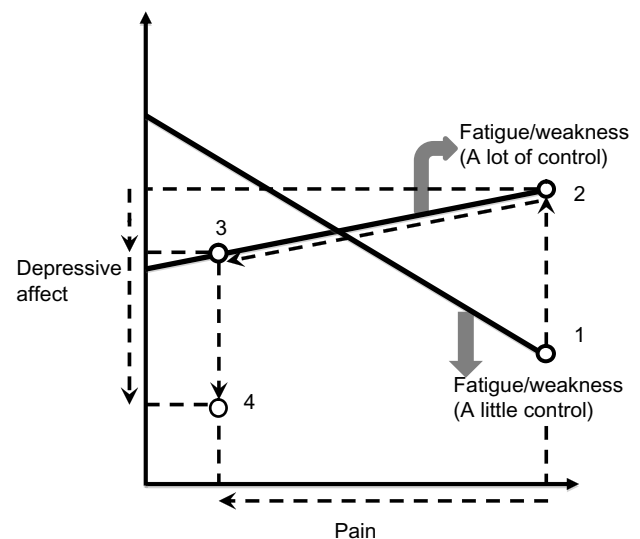

Figure 1 Potential context for a fatigue/weakness intervention with crossover impacts on pain and depressive affect.
Notes: In the current study, fatigue/weakness buffers the pain–depressive affect relationship, which implies that at a given level of pain, a countervailing intervention that only relieves fatigue/weakness would magnify the pain–depressive affect relationship (ie, moving along the dashed line from circled point 1 to 2). Furthermore, if the intervention also reduces pain, a lower level of depressive affect is predicted (ie, moving along the solid graphed line representing the magnifier effect, from circled point 2 to 3), although this resulting level of depressive affect is not necessarily equal to or lower than prior to adapting the intervention (rather, this resulting level of depressive affect depends on the relative magnitudes of the slopes from the two lines). In any event, a level of depressive affect at or below the original context may more likely be achieved if the intervention has a further and direct impact in reducing the level of the depressive affect outcome, beyond its indirect influence in relieving pain and fatigue/weakness (ie, moving along the dashed line from circled point 3 to 4). Thus, extrapolation from this graph suggests that interventions to relieve resistant fatigue/weakness in the absence of fever would also need to reduce pain and depressive affect in order to overcome the magnified pain–depressive affect relationship that would be predicted if only fatigue/weakness were relieved. Thus, Figure 1 illustrates how the original buffering effect may indicate that desirable crossover effects from a fatigue/weakness intervention could be achieved in a broader context where depression becomes another interacting symptom.

interventions may be plausible. Figure 1 illustrates how a buffering effect may be a signal that desirable crossover effects from a fatigue/weakness intervention could be achieved in a broader context. This graph of relationships from the current study suggests that interventions to relieve resistant fatigue/weakness in the absence of fever would also need to reduce pain and depressive affect as well in order to overcome the magnified pain–depressive affect relationship that would be predicted if only fatigue/weakness were relieved (note that using the study data, if the slope equation in Supplementary material [Part A], is estimated at "a little" and "a lot" of control of fatigue/weakness, the negative slope for the line representing a buffering effect is about twice as large in magnitude as the positive slope for the line representing a magnifier effect. The solid graphed lines in Figure 1 are drawn to reflect these specific choices).

The use of the psychostimulant methylphenidate, as an adjuvant to an opioid medication, qualifies as one intervention that may satisfy this particular context, since methylphenidate potentiates the effects of opioids in relieving pain[71,72] and also relieves fatigue and depression.[71] This context also suggests a plausible mechanism for how a cognitive-behavioral intervention could have reduced the severity of a pain, fatigue, and sleep disturbance cluster in a randomized controlled trial.[67] Based on post hoc findings derived in Part 1 of this article, I suggested that cognitive behavioral interventions could reduce depressive affect (an indicator of the physiological response of malaise during sickness behavior), which in turn mediates the reduced severity of the cluster symptoms.[3] Finally, to the extent that neuropathic pain is experienced, the complementary medicine approaches discussed earlier may be appropriate.

The reader is reminded that these patient profiles correspond to the initiation of palliative radiation in outpatients considered beyond cure. As the effects of palliative radiation become realized, there is the potential that this modality may relieve pain, resistant fatigue, and opioid side effects (sleepiness, fatigue, fever, subsequent pain). Moreover, the scope for using palliative radiation to relieve symptom clusters may be wider than commonly believed; a recent innovation uses curative or palliative radiation to stimulate the immune system, followed by immunotherapy to destroy remaining cancer cells.[73] This strategy could improve the attractiveness of starting palliative radiation earlier in the disease course when a curative focus still predominates.

Conclusion

Fever magnifies the pain–depressive affect relationship when sleep problems or fatigue/weakness are incompletely controlled. Thus, fever is supported as a biomarker/sign/symptom that aggravates the pain–depressive affect relationship in the context of the symptom cluster of pain-sleep problems-fatigue/weakness, suggesting a unique subgroup experiencing cytokine-mediated sickness behavior or analgesic side effects. Sleep problems and fatigue/weakness each magnify the pain–depressive affect relationship as well. These synchronistic comoderator effects compound the magnifier effects from fever, which suggests a promising context for crossover impacts by interventions for fever or fever mediators (anti-inflammatory cytokines, biomarkers, and pathways) that may relieve these other symptoms. For instance, research on symptom clusters should focus on crossover interventions targeting the cholinergic anti-inflammatory pathway, especially with patients who seek to reduce or quit smoking. However, it must also be stressed that fever is not always

harmful or distressing, and fever and its mediators should not be targeted for intervention without carefully considering the patient's individual clinical circumstances.

In participants without fever, occasional fatigue/weakness also magnifies the pain–depressive affect relationship and may provide a similar context for crossover impacts. However, more frequent fatigue/weakness buffers the pain–depressive affect relationship, suggesting other mechanism(s) and the need for crossover interventions with additional impacts that reduce depressive affect directly (and not only indirectly by relieving pain and fatigue/weakness). The existence of an appropriate context where crossover interventions could directly relieve all of these symptoms is suggested by a post hoc analysis of participants without fever in which fatigue/weakness, sleep problems, and depressive affect all serve to comagnify, across their full symptom ranges, the relationship between pain and mobility problems. The psychostimulant methylphenidate, in potentiating the effects of opioids in relieving pain[71,72] and in providing direct relief of fatigue and depression,[71] is one intervention that satisfies this context. Cognitive behavioral interventions might also satisfy this context.[2,3]

In fever and non-fever contexts, research on symptom clusters should confirm whether specific complementary medicine modalities yield crossover impacts when neuropathic pain presents. SRC could facilitate analysis of symptom clusters from these future investigations.

Methodological innovations can reveal novel approaches and lead to new kinds of opportunities that advance symptom cluster research. SRC holds promise to improve detection of statistical interactions among signs, symptoms, and/or biomarkers that can reveal causal and systemic complexity and improve translational research. In contrast with the call for multisite investigations of subgroup effects that cannot be detected in underpowered small samples (eg, Jacobsen and Jim[74]), the current study was feasible because SRC capitalizes on statistical power by eliminating inessential multicollinearity within a clinical sample of small to moderate size, thereby improving the capacity to detect subgroup effects. SRC cannot overcome, of course, sampling biases in small or non-representative samples, which limits the extent that current study findings can be generalized. However, in investigations of internal validity, SRC provides a new, efficient means to explore untapped population heterogeneity within targeted clinical samples in order to identify patient subgroups at heightened risk and contexts where interventions could relieve multiple symptoms. It could also foster expanded insight into interactions and synergistic relationships in many other areas of research.

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Disclosure

The author reports no conflicts of interest in this work.

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