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ORIGINAL RESEARCH

A common genetic variation in *CEBPE* and acute lymphoblastic leukemia: a meta-analysis of the available evidence

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Abstract: Acute lymphoblastic leukemia (ALL) has been studied intensively for decades, but the details of its etiology and underlying mechanisms have yet to be fully elucidated. It is now generally acknowledged that genetic factors contribute greatly to the development of this disease. The gene encoding CCAAT/enhancer-binding protein ε (*CEBPE*) is involved in the development of leukemia, and in particular the rs2239633 single nucleotide polymorphism (SNP) of CEBPE. The association between rs2239633 and risk of ALL has been well studied, but remains unclear. Therefore, a meta-analysis was performed in this study to establish a more precise estimation of that relationship. A comprehensive literature search of the PubMed electronic database was conducted, and relevant studies published up to February 20, 2015 were selected for analysis. The references of the retrieved articles were also screened. The extracted data were analyzed statistically, and pooled odds ratios with 95% confidence intervals were calculated using Review Manager (version 5.2) to estimate the association strength. Finally, eleven studies were included in the meta-analysis. The pooled analyses revealed that rs2239633 was associated with an increased risk of childhood ALL in Caucasians under any contrast models (P < 0.01). However, this SNP did not affect the risk of ALL in adulthood among Caucasians, or in childhood among East Asians. In conclusion, these findings confirm that the CEBPE rs2239633 SNP could be considered a good marker of pediatric ALL risk in Caucasians, but not in East Asians; it is not a good marker of adult ALL risk in Caucasians. Keywords: CEBPE, rs2239633, ALL, risk, meta-analysis

Background

Acute lymphoblastic leukemia (ALL) is the most common malignancy in childhood, accounting for 30% of all childhood cancers.¹ The etiology of childhood ALL is mostly unknown, although infections in the first years of life and some environmental factors, such as ionizing radiation and parental alcohol and tobacco use, are thought to play a causative role.²

ALL is known to result from an accumulation of mutations in tumor suppressor genes and oncogenes, and genetic alterations affecting several chromosomes.² Thus, the genetic variations may be important factors in its development. This finding has recently led to the performance of several genome-wide association studies to identify common single nucleotide polymorphisms (SNPs) affecting the susceptibility to childhood ALL.²⁻⁶ One of the most widely investigated SNPs identified in these studies is the common rs2239633 SNP of *CEBPE*, the gene encoding CCAAT/enhancer-binding protein ε (CEBPE), which plays an important role in the regulation of myelopoiesis.⁷ Papaemmanuil et al⁴ also found that CEBPE is a suppressor of myeloid leukemogenesis, and may be involved in B-cell precursor ALL development. The populations studied in the aforementioned genome-wide association studies were limited, and the many

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subsequent association studies that have been conducted with a view to verifying the described effects of rs2239633 on susceptibility to ALL in childhood in different populations have yielded varying results.^{8–13}

A meta-analysis was performed in the present study to review and compare all relevant studies, enabling more accurate conclusions to be drawn. In addition, since the genetic background of ALL differs between the adult and childhood forms, some studies were also performed to make clear if this SNP was a risk factor for adult ALL.^{14,15} In this study, those two relevant studies were also compared to examine the differential roles of rs2239633 in children and adults.

Materials and methods Literature search strategy

Two of the authors (Zhang and Du) each performed the literature search of the PubMed database independently. The process was supervised by the third author (Zhai), and any disagreement was resolved by consensus. The search strategy was to identify all possible studies that involved the use of the following test words: (14q11.2 "or" CEBPE "or" rs2239633) "and" (leukemia). The language of publication was restricted to English. The reference lists of retrieved articles were curated manually. The literature search was last updated on February 20, 2015. This study was judged exempt regarding ethics.

Selection criteria

All studies with the aim of investigating the association between presence of the rs2239633 SNP and ALL risk were potentially included. The following inclusion criteria were applied:

- 1. Published articles on human genetics (full texts or abstracts) without racial restriction.
- 2. If articles contained more than one geographic or other clinical characteristic subgroup, each subgroup was considered separately.
- 3. If multiple studies were derived from the same population, only the study with the largest sample was analyzed to avoid overlapping data.
- 4. Presentation of sufficient data on the distribution of rs2239633 in cases and in control groups or the data necessary to calculate those distributions. If the data were deficient or presented in an inappropriate form, the original authors were contacted and asked for the raw data.

Extracted information

A standard data-collection procedure was used, in line with the aforementioned inclusion criteria. Two of the authors (Zhang and Du) independently extracted the variables from the individual eligible studies in duplicate and made the characteristics compatible in a pooled database. Any encountered disagreements were resolved by discussion in order to reach a consensus. The following information was extracted from all eligible studies: name of the first author, year of publication, ethnicities of the study population, age and number of cases and controls, and main finding. The distribution of the genotypes among the controls was also tested for conformity to Hardy–Weinberg equilibrium (HWE).

Statistical analysis

The meta-analysis examined the overall association for allele contrasts, homozygote contrasts, and recessive and dominant models. The presence of heterogeneity was assessed using Cochran's *Q*-statistic (with a significance cutoff of P < 0.1) and quantified using the I^2 statistic, which is proportional to the degree of heterogeneity; the value of I^2 statistic lies in the range 0%–100%, with a value of 0% indicating homogeneity, and a value above 50% indicating the presence of a very high degree of heterogeneity.¹⁶

The raw data for genotype distribution were used to calculate the study-specific estimates of odds ratio (OR) and 95% confidence interval (CI). The overall pooled OR and 95% CI were estimated using the Mantel–Haenszel method, with a fixed-effects model when no significant clinical or statistical heterogeneity was present; when substantial heterogeneity was present, the random-effects model was used. The significance of the polled OR was determined by a test. The cutoff for statistical significance was set at P < 0.05.

Potential publication bias was estimated by constructing funnel plots. If most of the data appear at the top of a funnel plot and are distributed roughly symmetrically, this suggests the absence of obvious publication bias, and vice versa.¹³ There was no need to construct funnel plots when there were too few analyzed studies (ie, n < 5). All of these statistical analyses were performed using Review Manager (version 5.2) (Cochrane, London, UK).

The conformity of the rs2239633 SNP to HWE was tested among the controls in each individual study via the chi-square test or Fisher's exact test using SPSS software (SPSS Inc., Chicago, IL, USA). Sensitivity analyses were performed by excluding those studies for which the data deviated from HWE (P<0.05), omitting the highest weight or changing the model.

Results

Literature search and characteristics of the retrieved studies

Figure 1 presents a flow diagram summarizing the process of study selection. The final search generated 65 potentially



Figure I Flow chart of study selection process. Abbreviation: ALL, acute lymphoblastic leukemia.

relevant publications. After reading the titles and abstracts, 44 articles were found to be irrelevant to this study in that the research did not cover the association between rs2239633 and susceptibility to ALL, and were thus excluded from the analysis. Of the remaining 21 publications, 3 were case reports, 2 were reviews, and 1 was a conference proceedings; these were therefore also excluded. Finally, 15 publications were identified, in which the effect of rs2239633 on the risk

of both adult and childhood ALL was evaluated. Of these, four did not supply enough data, two were about adults,^{14,15} and nine were about children.^{4,8–13,17,18} These studies, providing enough data, were conducted in various populations of different ethnicities: eight involved Caucasians,^{4,10,11,13–15,17,18} two involved East Asians,^{9,12} and one involved multiple races.⁸ The basic data for every eligible study were extracted and are listed in Table 1. Although the authors of some of

Table I Study characteristics

Study	Ethnicity	Cases			Controls					
		Median age (year)	Number of males	Number of females	Median age (year)	Number of males	Number of females	HWE, P	results	
Emerenciano et al ⁸	Brazilian					262	243	0.135	N	
Lautner-Csorba et al ¹⁷	Caucasian	6.4±4.2	308	235	16.1±12.4	305	224	0.508	Ν	
Orsi et al ¹⁸	European	<15						0.546	Ν	
Papaemmanuil et al⁴	UK	5.5±3.3	322	255		725	713	0.779	Y	
Papaemmanuil et al ^{4,a}	UK	5.8±3.8	252	176				0.517	Y	
Pastorczak et al ¹¹	European	4.9	228	170	22	332	399	0.454	Ν	
Prasad et al ¹³	German	6	663	530	58	762	754	0.233	Y	
Prasad et al ^{13,a}	UK	6	101	90	59	117	244	0.310	Ν	
Ross et al ¹⁰	Caucasian							0.091	Ν	
Vijayakrishnan et al ¹²	Thai	6	109	81	18-25	82	100	0.162	Ν	
Wang et al ⁹	Chinese	1–18	349	219	1–18	441	231	0.147	Ν	
Peyrouze et al ¹⁵	European	18-60						0.355	Ν	
Burmeister et al ¹⁴	German	16–76	188	134				0.233	Ν	
Enciso-Mora et al ²⁴	UK	5.4±3.6	464	360					Y	
Migliorini et al ²⁵	UK/German								Y	
Chokkalingam et al ²²	Hispanic	5.6±3.6	156	144	5.5±3.6	214	192		Y	
Walsh et al ²³	Hispanic	5.2±3.3	156	141	5.3±3.4	240	214		Y	

Note: aIndicated another population studied in the same article.

Abbreviations: HWE, Hardy–Weinberg equilibrium; N, no; Y, yes.

the publications were contacted to obtain the raw genotype data, some of the data could not be obtained.

Meta-analysis

rs2239633 SNP and ALL risk in childhood

As stated, nine papers studied the association between rs2239633 and risk of ALL in childhood and provided data about genotype distribution. The data from these studies were pooled and subjected to a meta-analysis. Overall, this SNP was found to be significantly associated with risk of ALL in childhood under the dominant model (Figure 2A, P < 0.00001, OR =1.27, 95% CI =1.15–1.39), recessive model (Figure 2B, P < 0.00001, OR =1.23, 95% CI =1.14–1.33), and allele contrast (Figure 2C, P < 0.00001, OR =1.18, 95% CI =1.12–1.24). However, the presence of heterogeneity was found in the recessive model (Figure 2B, P=56%) and allele contrast (Figure 2C, P=57%), indicating the requirement for a subanalysis. Therefore, a submeta-analysis based on races was conducted. Among the included studies, the subjects in six were Caucasians,^{4,10,11,13,17,18} those

in two were East Asians,^{9,12} and those in the remaining study were a mixed ethnic population.⁸ The submetaanalysis revealed that under the three contrast models, rs2239633 was significantly associated with childhood ALL in Caucasians (Figure 2A, P < 0.00001, OR =1.31, 95% CI=1.18–1.46 for dominant model; Figure 2B, P < 0.00001, OR =1.32, 95% CI =1.21–1.44 for recessive model; Figure 2C, P < 0.00001, OR =1.23, 95% CI =1.16–1.30 for allele contrast), but not in East Asians (Figure 2). The degree of heterogeneity in both the Caucasian and East Asian subgroups was relatively small (Figure 2, P=28%, 29%, 36% for Caucasian under the dominant, recessive, and allele contrast models, respectively; P=0 for East Asian under the dominant, recessive, and allele contrast models.

rs2239633 and ALL risk in adults

Two of the studies investigated the relationship between the rs2239633 SNP and risk of ALL in adults, and provided the genotype data.^{14,15} The subjects in both of

Study or subgroup	All Events	Total	Control Events		Weight	Odds ratio M–H, fixed, 95%	CI	Odds ratio M–H, fixed, 95% Cl
Caucasians						,,		,,.
auther-Csorba et al ¹⁷	449	541	408	529	9.3%	1.45 (1.07, 1.96)		
Drsi et al ¹⁸	365	440	1,185	1,541	11.9%	1.46 (1.11, 1.93)		
Papaemmanuil et al4	425	503	1,103	1,435	11.8%	1.64 (1.25, 2.15)		
Papaemmanuil et al ^{4,a}	330	404	755	960	10.9%	1.21 (0.90, 1.63)		+
astorczak et al11	295	388	551	711	12.4%	0.92 (0.69, 1.23)		
Prasad et al13	996	1,193	1,203	1,510	23.2%	1.29 (1.06, 1.57)		
Prasad et al ^{13,a}	157	183	288	352	3.7%	1.34 (0.82, 2.20)		
loss et al ¹⁰	66	85	273	363	3.1%	1.15 (0.65, 2.01)		
Subtotal (95% CI)		3,737		7,401		1.31 (1.18, 1.46)		•
otal events	3,083		5,766					
leterogeneity: χ^2 =9.79,			3%					
est for overall effect: Z=	5.18 (<i>P</i> <0.0	0001)						
ast Asians								
'ijayakrishnan et al ¹²	179	190	173	182	1.4%	0.85 (0.34, 2.09)		
Vang et al ⁹	498	568	590	672	8.8%	0.99 (0.70, 1.39)		
Subtotal (95% CI)		758		854	10.2%	0.97 (0.71, 1.33)		-
otal events	677		763					
leterogeneity: $\chi^2 = 0.10$,	<i>df</i> =1 (<i>P</i> =0.7	5); /²=0%	%					
est for overall effect: Z=	0.19 (<i>P</i> =0.8	5)						
lixed	100	100	404	400	0.00/	0.07 (0.57.4.00)		
merenciano et al ⁸	139	160	421	483	3.6%	0.97 (0.57, 1.66)		
Subtotal (95% CI)		160		483	3.6%	0.97 (0.57, 1.66)		
otal events	139		421					
leterogeneity: not applic est for overall effect: Z=		2)						
otal (95% CI)		4,655		8,738	100%	1.27 (1.15, 1.39)		•
otal events	3,899		6,950					
leterogeneity: χ^2 =13.88	, df=10 (P=0).18); /²=	=28%				0.2	0.5 1 2

Figure 2 (Continued)

Study or subgroup	All Events	Total	Control Events	Total	Weight	Odds ratio M–H, fixed, 95% (Odds ratio CI M–H, fixed, 95% CI
Caucasians							
autner-Csorba et al17	172	541	152	529	9.1%	1.16 (0.89, 1.50)	+
Orsi et al ¹⁸	141	440	428	1,541	11.2%	1.23 (0.98, 1.54)	
Papaemmanuil et al⁴	181	503	381	1,435	11.0%	1.56 (1.25, 1.93)	
Papaemmanuil et al4.a	142	404	267	960	8.9%	1.41 (1.10, 1.80)	
Pastorczak et al11	119	388	207	711	8.8%	1.08 (0.82, 1.41)	
Prasad et al13	437	1,193	430	1,510	20.9%	1.45 (1.23, 1.71)	
Prasad et al ^{13,a}	62	183	105	352	4.1%	1.21 (0.82, 1.77)	
Ross et al ¹⁰	23	85	108	363	2.6%	0.88 (0.52, 1.49)	
Subtotal (95% CI)		3,737		7,401	76.8%	1.32 (1.21, 1.44)	•
Total events	1,277		2,078				
Heterogeneity: $\chi^2=9.87$	7, df=7 (P=	0.20); /2=2	29%				
Test for overall effect: 2	Z=6.25 (P<	0.00001)					
East Asians							
/ijayakrishnan et al12	103	190	95	182	3.9%	1.08 (0.72, 1.63)	
Vang et al ⁹	245	568	309	672	14.0%	0.89 (0.71, 1.12)	+
Subtotal (95% CI)		758		854	17.9%	0.93 (0.77, 1.14)	-
Total events	348		404				-
Heterogeneity: $\chi^2=0.68$ Test for overall effect: 2			0%				
Mixed							
Emerenciano et al ⁸	71	160	220	483	5.3%	0.95 (0.67, 1.37)	
Subtotal (95% CI)		160		483	5.3%	0.95 (0.67, 1.37)	-
Fotal events Heterogeneity: not app	71 licable		220				
Test for overall effect: 2	Z=0.26 (P=	0.80)					
Fotal (95% CI)		4,655		8,738	100%	1.23 (1.14, 1.33)	
			0.700			,	1
Total events	1,696		2,702				

Test for subgroup difference: χ^2 =11.93, *df*=2 (*P*=0.003); *I*²=83.2%

Study or subgroup	All Events	Total	Control Events	Total	Weight	Odds ratio M–H, fixed, 95% (Odds ratio	
Caucasians								
Lautner-Csorba et al ¹⁷	621	1,082	560	1,058	9.4%	1.20 (1.01, 1.42)	-	-
Orsi et al ¹⁸	506	880	1,613	3,082	11.8%	1.23 (1.06, 1.43)		-
Papaemmanuil et al⁴	606	1,006	1,484	2,870	11.9%	1.41 (1.22, 1.64)		
Papaemmanuil et al4,a	472	808	1,022	1,920	9.8%	1.23 (1.05, 1.46)		
Pastorczak et al11	414	776	758	1,422	9.7%	1.00 (0.84, 1.19)	-	_
Prasad et al13	1,433	2,386	1,633	3,020	22.4%	1.28 (1.15, 1.42)		-e-
Prasad et al ^{13,a}	219	366	393	704	4.2%	1.18 (0.91, 1.52)	+	
Ross et al ¹⁰	89	170	381	726	2.7%	0.99 (0.71, 1.39)	-	
Subtotal (95% CI)		7,474		14,802	81.9%	1.23 (1.16, 1.30)		+
Total events	4,360		7,844					
Heterogeneity: $\chi^2 = 11.0$,	df=7 (P=0.1	14); /²=36	5%					
Test for overall effect: Z=	=7.05 (P<0.0	00001)						
East Asians								
Vijavakrishnan et al ¹²	282	380	268	364	2.7%	1.03 (0.74, 1.43)		
Wang et al ⁹	743	1,136	899	1,344	11.1%	0.94 (0.79, 1.11)	_	-
Subtotal (95% CI)		1,516		1,708	13.8%	0.95 (0.82, 1.11)		
Total events	1,025		1,167					
Heterogeneity: $\chi^2=0.27$, Test for overall effect: Z=			%					
Mixed								
Emerenciano et al ⁸	210	320	641	966	4.3%	0.97 (0.74, 1.26)	_	_
Subtotal (95% CI)		320		966	4.3%	0.97 (0.74, 1.26)		•
Total events Heterogeneity: not app	210 licable		641					
Test for overall effect: Z	=0.24 (<i>P</i> =0.8	31)						
Total (95% CI)		9,310		17,476	100%	1.18 (1.12, 1.24)		•
Total events	5,595		9,652					
Heterogeneity: $\chi^2 = 23.14$	l, df=10 (P=	0.01); <i>I</i> 2:	=57%					
Test for overall effect: Z:							0.2 0.5 1	2

Figure 2 Meta-analysis of the association between rs2239633 and childhood ALL risk.

Notes: (A) Under the dominant model; (B) under the recessive model; (C) under the allele contrast model. Indicated another population studied in the same article. Abbreviations: ALL, acute lymphoblastic leukemia; CI, confidence interval; M–H, Mantel–Haenszel method.

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those studies were Caucasians. After pooling these data, further analysis revealed that this SNP was not associated with risk of ALL in adults in any model of comparison (Figure 3A, P=0.18, OR =1.21, 95% CI =0.91–1.61 for dominant model; Figure 3B, P=0.79, OR =1.03, 95% CI =0.81–1.32 for recessive model; Figure 3C, P=0.35, OR =1.08, 95% CI =0.92–1.26 for allele contrast). There was no heterogeneity in any of the three contrast models (Figure 3, P=0).

Publication bias and sensitivity analysis

The publication bias for the research with pediatric subjects was assessed. Overall, no obvious publication bias was found under the recessive model, and the funnel plot exhibited good symmetry under the dominant model if the study of Vijayakrishnan et al¹² was excluded. The result did not change after excluding this article. The symmetry of the funnel plot was poor under the allele contrast model, which suggested obvious publication bias (Figure 4).

The stability of the results involving Caucasians was evaluated by performing a sensitivity analysis. Neither omitting Prasad et al's¹³ study (highest weight) nor changing the model (random-effects model) made a significant difference to the pooled effects of the three contrast models, suggesting that the meta-analysis findings were highly stable.

Discussion

CEBPE is involved in terminal differentiation and functional maturation of neutrophils and macrophages, and is critical for transcription of most granule proteins localized to specific and gelatinase granules as well as for azurophil granule proteins expressed in the late promyelocyte stage.^{19,20} It has also been reported that CEBPE is targeted by recurrent immunoglobulin heavy-chain (locating on chromosome band 14q32) chromosomal translocations in B-cell precursor ALL, and mutations in *CEBPE* lead to neutrophil-specific granule deficiency.^{9,19,20} All of these findings suggested that *CEBPE* plays a role in the development of ALL. The studies

Α	Study or subgroup	All Events	Total	Control Events	Total	Weight	Odds ratio M–H, fixed, 95% Cl			ratio fixed, 95	5% CI	
	Burmeister et al ¹⁴ Peyrouze et al ¹⁵	265 80	322 100	1,203 84	1,510 112	82.5% 17.5%	1.19 (0.87, 1.62) 1.33 (0.70, 2.56)					
	Total (95% CI)		422		1,622	100%	1.21 (0.91, 1.61)			•		
	Total events	345		1,287								
	Heterogeneity: χ^2 =0.10, Test for overall effect: Z=			1%				0.01	0.1	1	10	- I 100
в	Study or subgroup	All Events	Total	Control Events	Total	Weight	Odds ratio M–H, fixed, 95% Cl			ratio fixed, 95	5% CI	
	Burmeister et al14	92	322	430	1,510	83.8%	1.00 (0.77, 1.31)					
	Peyrouze et al ¹⁵	33	100	33	112	16.2%	1.18 (0.66, 2.11)			- - -		
	Total (95% CI)		422		1,622	100%	1.03 (0.81, 1.32)			•		
	Total events	125		463								
	Heterogeneity: χ^2 =0.24, Test for overall effect: Z=	•		1%				0.01	0.1	1	10	100
С	Study or subgroup	All		Control		Weight	Odds ratio		Odds	ratio		

U	Study or subgroup	y or subgroup All		Control	Control		Odds ratio	Odds ratio				
		Events	Total	Events	Total		M–H, fixed, 95% Cl		М-Н,	fixed, 95	% CI	
	Burmeister et al14	357	644	1,633	3,020	84.2%	1.06 (0.89, 1.25)					
	Peyrouze et al ¹⁵	113	200	117	224	15.8%	1.19 (0.81, 1.74)			-		
	Total (95% CI)		844		3,244	100%	1.08 (0.92, 1.26)			•		
	Total events	470		1,750								
	Heterogeneity: $\chi^2=0.30$,	<i>df</i> =1 (<i>P</i> =0.	58); /²=0	1%				0.01	0.1	1	10	100

Test for overall effect: Z=0.93 (P=0.35)

Figure 3 Meta-analysis of the association between rs2239633 and adult ALL risk.

Notes: (A) Under the dominant model; (B) under the recessive model; (C) under the allele contrast model. Abbreviations: ALL, acute lymphoblastic leukemia; CI, confidence interval; M–H, Mantel–Haenszel method.



Figure 4 The publication bias for researches of childhood.

Notes: (A) Under the dominant model; (B) under the recessive model; (C) under the allele contrast model. Abbreviation: SE, standard error.

conducted thus far on the relationship between rs2239633 and risk of ALL have yielded a relatively small amount of data and have produced conflicting results. Those data were evaluated in the present meta-analysis. The primary findings of this investigation are the increased OR of childhood ALL susceptibility associated with presence of the rs2239633 SNP in a fixed-effects model, and the lack of an association between this SNP and risk of ALL in adults.

Subgroup analysis for childhood ALL according to race revealed that this SNP impacted Caucasians, and not Asians. This finding was extremely robust, remaining significant irrespective of the different model assumptions (fixedeffects versus random-effects models), genetic models, and additional sensitivity analysis. This disparity between races may reflect the existence of a true population-specific disease variant, but may also be attributable to differences in genomic structure at these loci between populations. Wang et al⁹ determined the linkage disequilibrium (LD) patterns at the locus of rs2239633 in European and Chinese populations from the HapMap database, and found that rs2239633 is located in different LD blocks in the two populations. Therefore, we speculate that the pathogenic locus might not be rs2239633, but rather other SNPs in the LD block. Other factors contributing to the differential results for the two populations include differences in allele frequency and differences in genetic and environmental backgrounds that interact with the variants.¹²

It has been proposed that genetics plays a role in adult cancer susceptibility, although this has been difficult to evaluate because of confounders, such as environmental factors, the immune system, and exposure to pathogens. To the best of our knowledge, no predisposing genetic marker has been proven for adult ALL. No association between rs2239633 and risk of adult ALL was found in this study. Although the failure to detect any weak association may be affected by the relatively small sample, these data suggest that the genetic impact of polymorphisms at these loci is at least weaker for adult ALL than it is for pediatric ALL and highlights genetic differences between ALL occurring in adults versus children.¹⁵

Based on a meta-analysis, Wang et al²¹ concluded that rs2239633 confers an increased risk of childhood ALL, especially among Caucasians and Hispanics. The main findings of the present meta-analysis are consistent with that conclusion.

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While, the adults' ALL risk related with rs2239633 was also evaluated, three different genetic models were used in our study, which showed more comprehensive evaluation of the impact of rs2239633 on ALL. In addition, we believe that the article of Wang et al²¹ contained some data errors. The research data – and in particular the OR and 95% CI values - of Chokkalingam et al²² and Walsh et al²³ that were used in the meta-analysis performed by Wang et al²¹ were calculated using logistic regression. We speculate that since the OR and 95% CI values generally differed before and after data adjustment, Wang et al²¹ did not use the raw data from the studies of Chokkalingam et al²² and Walsh et al²³ thus potentially causing bias. It is not possible to derive raw data using the adjusted data. Before completing our study, we attempted to contact the authors of those two articles to obtain the raw genotype data,^{22,23} but were unsuccessful. Therefore, these two articles were excluded from the analysis, which we believe was the most appropriate action.

It is worth noting that there were no differences in the results among the three contrast models, although the funnel figure obtained under the allele contrast model exhibited poor symmetry, and a clear outlier appeared in that obtained under the dominant model, indicating that the impact of genetic models on the results is not large.

Our study demonstrated that genotype testing is very important, although just a statistical association was found. Based on genotype testing, the susceptible population will be able to take the necessary protective measures, which is particularly important for those with long-term exposure to environmental pollutants.

This study had some limitations. First, the effect of gene–gene and gene–environmental interactions was not addressed in the meta-analysis due to lack of data. Second, the meta-analysis was based on unadjusted estimates, and the confounding factors could not be controlled for because most studies did not provide these data. Third, for publication bias, the best symmetry was observed for the funnel figure produced under the recessive genetic model, but not under the dominant model and allele contrast. Finally, the conclusions regarding risk of ALL among adults and Asian children are based on relatively few cases, and so the statistical power thereof is not strong.

Conclusion

In summary, the findings of the present meta-analysis support a positive association between *CEBPE* rs2239633 and childhood risk of ALL in Caucasians, but there was no such association for risk of ALL in Caucasian adults or East Asian

children. Further large-scale studies assessing gene–gene and gene–environment interactions are required.

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Disclosure

The authors report no conflicts of interest in this work.

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