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ORIGINAL RESEARCH

Genetic variants in insulin-like growth factor binding protein-3 are associated with prostate cancer susceptibility in Eastern Chinese Han men

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Background: Growing evidence has indicated that insulin-like growth factor binding protein-3 (IGFBP-3) polymorphisms are associated with altered risk of prostate cancer (PCa). However, few studies have been conducted in Chinese population to validate this association.

Materials and methods: Herein, we examined the association between genetic variants in the IGFBP-3 gene and PCa risk in the Chinese Han population based on a genome-wide association study (1,417 cases and 1,008 controls), and replicated three genetic variants loci in an independent case-control study (1,755 cases and 1,523 controls) using Sequenom platform. Logistic regression analyses were performed to estimate odds ratios (ORs) and 95% confidence intervals (95% CIs).

Results: We found that in the discovery stage, rs9691259 (OR =0.691, 95% CI: 0.587–0.814, P<0.001) and rs6950179 (OR =1.420, 95% CI: 1.201-1.677, P<0.001) were significantly associated with PCa risk, whereas rs2854744 showed a marginal association with PCa risk. In the replication stage, the association between rs9691259 and rs6950179 and PCa risk was not replicated, whereas rs2854744 conferred a significant association with PCa risk (OR =1.399, 95% CI: 1.010–1.937, P=0.043). After combining the two stages, we found that rs9691259, rs6950179, and rs2854744 were all significantly associated with PCa risk.

Conclusion: This study suggests that IGFBP-3 genetic variants are significantly associated with PCa risk in the Chinese population.

Keywords: IGFBP-3, polymorphism, case-control study, genetic susceptibility

Introduction

Over the past few decades, much attention has been paid to the important role of metabolic syndrome (MS) in carcinogenesis. As a key factor of MS, insulin resistant denotes an interrupted state in the biological response to insulin. Under these circumstances, disordered expression and/or function of the insulin-like growth factor (IGF) axis is considered to be involved in the development and progression of various cancers, including prostate cancer (PCa). IGF binding protein-3 (IGFBP-3) is an important member of the IGFBP family, which binds more than 90% of circulating IGF-1 with high affinity and specificity, and regulates its biological activities. In vitro studies have demonstrated that IGFBP-3 independently inhibits cell proliferation and induces apoptosis of PCa cells.^{1,2} In addition, IGFBP-3 can exert an inhibitory role in PCa cell adhesion and motility.³ Moreover, using IGFBP-3 knockout mice, Mehta et al revealed that IGFBP-3 can suppress distant PCa metastases.⁴

Although PCa incidence is significantly lower in the People's Republic of China than in Western countries, disease-related morbidity is increasing year by year according

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to recent epidemiological studies.5 Genetic influence that explains 57% of all PCa based on the results of a twin study is considered vital;⁶ however, PCa etiology is still not well determined. Some rare germline mutations, such as HOXB13, were demonstrated to be closely associated with PCa risk in Chinese men,⁷ whereas germline sequence variants of some other important genes participating in PCa development, PTEN for instance, were not found to be intimately linked with PCa susceptibility.8 Likewise, the impact of genetic variants in IGFBP-3 on PCa risk varies significantly. Hernandez et al reported a two-fold increase in PCa risk for African-American men homozygous for the IGFBP-3 -202 C allele (rs2854744).9 Another study group observed no significant associations between IGFBP-3 rs2854744 and PCa risk in both African-Americans and Caucasians.¹⁰ These inconsistent results may embody the influence of different ethnic backgrounds and environmental factors. Furthermore, most of these studies are limited to developed countries or a few single-nucleotide polymorphisms (SNPs), and no study to date has been conducted to examine the relationship between IGFBP-3 genetic variants and PCa risk in the Chinese population. Therefore, more polymorphic sites from a large-scale population are important to better understand the role of IGFBP-3 SNPs in PCa carcinogenesis. In the present study, we performed, for the first time in Southeastern Chinese men, a two-stage casecontrol study to evaluate whether IGFBP-3 polymorphisms are associated with PCa risk.

Materials and methods Study subjects

The present case-control study included two stages: data from a genome-wide association studies (GWAS) stage and a replication stage. The recruitment of study subjects in the GWAS stage has been described previously.¹¹ Briefly, there were 1,417 hospital-based cases with pathologically confirmed PCa and 1,008 cancer-free controls that came from four districts and counties in the Shanghai, People's Republic of China area in the discovery stage. The replication stage included 1,755 pathologically confirmed PCa cases and 1,523 controls that visited for routine physical examination in local hospitals in Shanghai, Nanjing, and surrounding areas. All subjects were unrelated Han Chinese, and cases and controls were not overlapped in the two stages. Blood samples were obtained from each participant for DNA extraction once consent was acquired. Essential information, such as age, prostate-specific antigen (PSA) levels, history of cancer, as well as pathological data of PCa, was collected from standardized questionnaires and medical records. The study protocols were approved by the Institutional Research Review Boards of Fudan University Shanghai Cancer Center, Huashan Hospital, Fudan University, and other participating institutions, and written informed consent was obtained from all subjects before participation.

SNP genotyping

All SNPs in the IGFBP-3 gene were selected from the International HapMap Project database (http://hapmap.ncbi. <u>nlm.nih.gov/</u>) for genotyping in the discovery stage, and samples were genotyped on the MassARRAY iPLEX (Sequenom Inc., San Diego, CA, USA) platform. Standard quality control procedures to filter out unqualified DNA samples and SNPs were performed as described previously.Untyped SNPs were input according to reference data from the 1,000 Genomes Project CHB + JPT subjects by the IMPUTE2 program.¹² Based on the following criteria, qualified SNPs were filtered as follows: 1) imputation quality score >0.8; 2) call rate >95%; 3) minor allele frequency $\geq 5\%$ in Chinese Han populations; and 4) P-value for Hardy-Weinberg equilibrium in controls >0.001. In the replication stage, genotyping was performed using the Sequenom platform, and allelic discrimination was analyzed by the SDS v 2.4 (Applied Biosystems, Inc., Carlsbad, CA, USA) software. Two negative controls were included in each reaction plate, and genotyping was randomly repeated in 10% of samples, achieving a 100% concordance for all samples. Technicians who performed the genotyping procedures were blinded to the status of all subjects.

Statistical methods

Pearson's chi-squared tests were performed to analyze the agreement with Hardy–Weinberg equilibrium among the controls. Unconditional logistic regression analysis was used to estimate odds ratios (ORs) and 95% confidence intervals (CIs) for PCa risk with adjustment for age. All statistical tests were two-sided, and carried out by the SAS software (version 9.2; SAS Institute Inc., Cary, NC, USA).

Results

The clinicopathological characteristics of all subjects are listed in Table 1. There was no significant difference in age between cases and controls, in either the discovery stage or the replication stage. As expected, PCa patients presented with much higher PSA levels compared with controls. In both stages, the Gleason scores of patients were mostly \geq 7.

We used data from the GWAS stage where a total of 690 SNPs in the IGFBP-3 gene were detected. Nine SNPs (rs9691259, rs6964415, rs7793447, rs17172664, rs6976880,

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| Variables | Discovery stage | | Replication stage | |
|---------------------------------|-----------------|------------|-------------------|--------------|
| | Case | Control | Case | Control |
| | (n=1,417) | (n=1,008) | (n=1,755) | (n=1,523) |
| Ageª (mean ± SD) | 71.3±8.1 | 62.1±10.0 | 70.1±7.7 | 67.9±6.3 |
| PSA levels [♭] (ng/mL) | | | | |
| ≤4.0 | 55 (3.9) | 965 (95.8) | 32 (1.8) | 1,373 (90.2) |
| >4.0, <10.0 | 186 (13.1) | 32 (3.2) | 275 (15.6) | 118 (7.7) |
| ≥10.0 | 1,098 (77.5) | 9 (0.9) | 968 (55.2) | 32 (2.1) |
| Missing | 78 (5.5) | 2 (0.1) | 480 (27.4) | 0 (0) |
| T stage | | | | |
| TI | 180 (12.7) | NA | 79 (4.5) | NA |
| T2 | 547 (38.6) | NA | 412 (23.5) | NA |
| Т3 | 359 (25.3) | NA | 224 (12.8) | NA |
| T4 | 137 (9.7) | NA | 186 (10.6) | NA |
| Missing | 194 (13.7) | NA | 854 (48.6) | NA |
| N stage | | | | |
| N0 | 786 (55.5) | NA | 740 (42.2) | NA |
| NI | 369 (26.0) | NA | 144 (8.2) | NA |
| Nx | 262 (18.5) | NA | 871 (49.6) | NA |
| M stage | | | | |
| M0 | 832 (58.7) | NA | 680 (38.8) | NA |
| MI | 440 (31.1) | NA | 234 (13.3) | NA |
| Mx | 145 (10.2) | NA | 841 (47.9) | NA |
| Gleason score | | | | |
| <7 | 355 (25.1) | NA | 243 (13.8) | NA |
| ≥7 | 993 (70.1) | NA | 895 (51.0) | NA |
| Missing | 69 (4.8) | NA | 617 (35.2) | NA |

Table I Clinicopathological characteristics of subjects in the two stages

Notes: Values are presented as n(%). *Age at diagnosis for cases and at recruitment for controls. *PSA levels were determined at diagnosis for cases and at recruitment for controls.

Abbreviations: NA, not available; PSA, prostate-specific antigen; SD, standard deviation.

rs6950179, rs1077750, rs7794166, and rs6970565) achieved a statistical level of 10^{-5} . Among these nine SNPs, rs6964415, rs7793447, rs17172664, and rs6976880 were in strong linkage disequilibrium (LD; $r^2>0.8$) with rs9691259, whereas rs1077750, rs7794166, and rs6970565 were in strong LD ($r^2>0.8$) with rs6950179. Among the two groups, rs9691259 and rs6950179 reached the highest statistical level, respectively. Hence, we selected these two SNPs for further validation. In addition, rs2854744 indicated a marginal association with PCa risk (P=0.052). Because quite a few studies to date have reported that rs2854744 is associated with the risk of developing many cancers, $^{9,13-15}$ including PCa, we included this SNP in our replication stage.

Detailed information about the relationship between the three SNPs and PCa risk is indicated in Table 2. In the discovery stage, the G allele of rs9691259 was associated with a decreased risk of PCa, whereas the C allele of rs6950179 was associated with an increased PCa risk. The A allele of rs2854744 conferred a marginal association with an increased PCa risk. In the replication stage, the association between rs9691259 and rs6950179 and PCa risk was not replicated.

However, patients carrying the homozygous mutant of rs2854744 presented with an increased PCa risk after adjustment for age (OR =1.399, 95% CI: 1.010–1.937, P=0.043). A significant association was also observed in the dominant genetic model regarding rs2854744 and PCa risk (OR =1.168, 95% CI: 1.013–1.347, P=0.032). After combining data from the two stages, rs9691259, rs6950179, and rs2854744 were all significantly associated with PCa risk (Table 2).

Discussion

In the present two-stage case-control study, we found that three IGFBP-3 genetic variants (rs9691259, rs6950179, and rs2854744) were significantly associated with PCa risk. To the best of our knowledge, this is the first report of an association between IGFBP-3 polymorphisms and PCa risk in the Chinese Han population.

The incidence of MS is increasing worldwide, and MS-related cancers, including PCa, have become a public health issue. Lund Haheim et al conducted a large-scale research project in Scandinavia and reported that MS was positively associated with PCa risk.¹⁶ Another study found

| see s | HWF Cases Control | | | Replicat | Replication stage | ge | | | Overall | _ | | |
|--|-------------------|-----------------------------|-------|----------|--------------------------|----------|---------------------|-------|----------|----------|---------------------|-------|
| 0.639 842 507 1.00 487 410 0.715 (0.603-0.849) 86 89 0.582 (0.424-0.798) 0.691 (0.587-0.814) 0.667 (0.490-0.907) 0.088 497 436 1.00 0.088 497 436 1.00 0.088 497 436 1.00 524 129 1.523 (1.184-1.960) 1.420 (1.201-1.677) 1.777 1.292 694 1.00 0.174 922 694 1.00 0.174 922 694 1.00 | | Is OR ^a (95% CI) | ٩ | НМЕ | Cases | Controls | ORª (95% CI) | ٩ | Cases | Controls | ORª (95% CI) | ٩ |
| 0.639 842 507 1.00 487 410 0.715 (0.603-0.849) 86 89 0.582 (0.424-0.798) 0.691 (0.587-0.814) 0.667 (0.490-0.907) 0.088 497 436 1.00 0.088 497 436 1.00 666 421 1.388 (1.162-1.657) 666 421 1.388 (1.162-1.657) 1.224 129 1.523 (1.184-1.960) 1.420 (1.201-1.677) 1.777 0.174 922 694 1.00 0.174 922 694 1.00 | | | | | | | | | | | | |
| 487 410 0.715 (0.603-0.849) vs AA 86 89 0.582 (0.424-0.798) 0.691 (0.587-0.814) 0.667 (0.490-0.907) 0.108 497 436 1.00 0.088 497 436 1.00 0.088 497 436 1.00 0.088 497 436 1.00 0.088 497 436 1.00 0.174 922 694 1.00 0.174 922 694 1.00 0.174 922 694 1.00 | 842 | 1.00 | | 0.237 | 552 | 838 | 1.00 | | 1,394 | I,345 | 00.1 | |
| 86 89 0.582 (0.424-0.798) vs AA 0.691 (0.587-0.814) 0.667 (0.490-0.907) A/AG 0.667 (0.490-0.907) 0.667 (0.490-0.907) A/AG 0.0108 497 436 1.00 0.088 497 436 1.00 1.420 (1.62-1.657) 0.088 497 436 1.00 1.420 (1.201-1.677) vs TT 129 1.523 (1.184-1.960) 1.420 (1.201-1.677) XTT 0.174 922 694 1.00 0.174 922 694 1.00 1.152 (0.965-1.376) | | 0.715 (0.603-0.849) | | | 367 | 559 | 0.997 (0.841–1.181) | 0.997 | 854 | 969 | 0.850 (0.755-0.957) | 0.007 |
| Vs AA A/AG 0.667 (0.490-0.907) 0.667 (0.490-0.907) 0.667 (0.490-0.907) 666 421 1.00 1.388 (1.162-1.657) 224 129 1.523 (1.184-1.960) 1.420 (1.201-1.677) 1.777 0.174 922 694 1.00 0.174 922 694 1.00 | | 0.582 (0.424–0.798) | | | 68 | 109 | 0.947 (0.687–1.306) | 0.740 | 154 | 198 | 0.750 (0.600-0.938) | 0.012 |
| A/AG 0.667 (0.490-0.907) 0.088 497 436 1.00 666 421 1.388 (1.162-1.657) 224 129 1.523 (1.184-1.960) 1.420 (1.201-1.677) 1.280 (1.013-1.617) 0.174 922 694 1.00 0.174 922 694 1.00 0.174 922 694 1.00 | | 0.691 (0.587–0.814) | | | | | 0.989 (0.841–1.162) | 0.889 | | | 0.833 (0.745-0.933) | 0.002 |
| 0.088 497 436 1.00 666 421 1.388 (1.162–1.657) 224 129 1.523 (1.184–1.960) 1.420 (1.201–1.677) 1.777 1.280 (1.013–1.617) 0.174 922 694 1.00 | | 0.667 (0.490–0.907) | | | | | 0.948 (0.693–1.298) | 0.741 | | | 0.801 (0.644-0.996) | 0.046 |
| 0.088 497 436 1.00 666 421 1.388 (1.162–1.657) 224 129 1.523 (1.184–1.960) 1.420 (1.201–1.677) 1.71T 1.280 (1.013–1.617) 0.174 922 694 1.00 0.174 922 694 1.00 | | | | | | | | | | | | |
| 666 421 I.388 (I.162–I.657) 224 129 I.523 (I.184–I.960) . 3 TT 1.420 (I.201–I.677) 1.71T I.280 (I.013–I.617) 0.174 922 694 1.00 447 292 1.152 (0.965–I.376) | 497 | 1.00 | | 0.233 | 644 | 586 | 1.00 | | I, I 4 I | 1,022 | 00.1 | |
| 224 129 1.523 (1.184-1.960) AS TT 224 129 1.523 (1.184-1.960) T/TT 1.420 (1.201-1.677) 1.280 (1.013-1.617) 0.174 922 694 1.00 447 292 1.152 (0.965-1.376) 447 292 1.152 (0.965-1.376) | | I.388 (I.162–1.657) | | | 847 | 726 | 1.062 (0.914–1.233) | 0.433 | 1,513 | I, I 47 | 1.182 (1.054–1.324) | 0.004 |
| <pre>%TT 1.420 (1.201-1.677) T/TT 1.280 (1.013-1.617) 0.174 922 694 1.00 447 292 1.152 (0.965-1.376) 447 292 1.152 (0.965-1.376)</pre> | | 1.523 (1.184–1.960) | | | 243 | 197 | 1.122 (0.901–1.397) | 0.301 | 467 | 326 | 1.283 (1.088–1.513) | 0.003 |
| T/TT 1.280 (1.013-1.617) 0.174 922 694 1.00 447 292 1.152 (0.965-1.376) | | 1.420 (1.201–1.677) | | | | | 1.075 (0.932–1.239) | 0.321 | | | 1.204 (1.081–1.341) | 0.001 |
| 0.174 922 694 1.00 447 292 1.152 (0.965–1.376) | | 1.280 (1.013–1.617) | | | | | 1.085 (0.887–1.328) | 0.427 | | | I.I7I (I.005–I.363) | 0.043 |
| 922 694 1.00 447 292 1.152 (0.965–1.376) | | | | | | | | | | | | |
| 292 I.152 (0.965–I.376) | 922 | 1.00 | | 0.909 | I,004 | 922 | 00.1 | | 1,926 | 1,616 | 00.1 | |
| | | 1.152 (0.965–1.376) | 0.117 | | 612 | 494 | 1.138 (0.981–1.320) | 0.089 | 1,059 | 786 | I.130 (I.009–I.266) | 0.034 |
| (0,782–2.74b) 1.042 (0,782–2.74b) | 48 22 | 1.642 (0.982–2.746) | 0.059 | | 66 | 65 | 1.399 (1.010-1.937) | 0.043 | 147 | 87 | 1.418 (1.079–1.863) | 0.012 |
| | | 1.187 (0.999–1.410) | 0.052 | | | | 1.168 (1.013–1.347) | 0.032 | | | 1.159 (1.039–1.293) | 0.008 |
| AA vs AC/CC 1.571 (0.942–2.620) 0.083 | | 1.571 (0.942–2.620) | 0.083 | | | | 1.335 (0.968–1.840) | 0.078 | | | I.360 (I.037–I.782) | 0.026 |

CL, 95% confidence interval; HWE, Hardy–Weinberg equilibrium; IGFBP-3, insulin-like growth factor binding protein-3; SNP, single-nucleotide polymorphism. Abbreviations: OR, odds ratio; 95%

that MS was closely related with PCa progression.¹⁷ As the key mechanism responsible for MS, IGFs and IGFBPs are thought to be important in PCa carcinogenesis. However, despite the inhibitory role of IGFBP-3 in vitro in PCa development, conflicting conclusions have been reported regarding the relationship between serum IGFBP-3 levels and PCa risk, as well as IGFBP-3 polymorphisms and PCa risk. The association between decreased serum IGFBP-3 levels and increased PCa risk has been reported,^{18,19} whereas some studies did not observe an influence of IGFBP-3 levels on PCa risk.^{20,21} In a multi-ethnic large-scale study including 2,320 PCa cases and 2,290 controls, Cheng et al found that common genetic variation in the IGFBP-3 gene did not substantially influence PCa susceptibility.¹³ Similar results were also reported by other studies, although some found a significant association between IGFBP-3 genetic variants and serum IGFBP-3 concentration.^{10,15} However, significant relationship between IGFBP-3 polymorphisms and PCa risk has been reported by many studies.9,18 Likewise, our study observed a positive association between IGFBP-3 genetic variants and PCa risk. Furthermore, Hoyo et al found that the association between IGFBP-3 genetic variants and PCa susceptibility was only significant in those with Gleason score <7 rather than ≥ 7 , in both white and black men.²² There are several explanations for these inconsistent results. As we know, genetic susceptibility may differ according to ethnic background, and natural selection can result in genetic diversity that might be sensitive to local environments. In addition, many factors, such as insufficient statistical power, differences in study population characteristics, allelic heterogeneity, allele frequency and/or LD differences across various races, population stratification, differences in effect sizes, and gene-environment interactions, may contribute to the lack of replication of genetic associations or the observation of different relationships.23 Therefore, further studies combining more ethnic backgrounds are necessary to validate our observations.

Previous studies regarding the association between IGFBP-3 genetic variants and cancer mainly focused on a specific SNP, rs2854744, which locates at position -202 in the promoter region of IGFBP-3 gene. This A/C polymorphism has been reported to be highly correlated to circulating IGFBP-3 levels, and mutant C allele is associated with decreased IGFBP-3 levels.14,24 This may explain in part the underlying mechanisms of the impact of genetic variants on clinical phenotype. Unfortunately, serum IGFBP-3 levels were not determined in our study, and we could not assess whether IGFBP-3 polymorphisms contribute to circulating IGFBP-3 levels. However, the present study has several strengths. First, we performed our research with a relatively large number of samples in the People's Republic of China, which may enhance our statistical power. Second, we validated our observations in a replication stage based on a GWAS dataset. Compared with most previous studies that examined limited SNPs, our GWAS stage covered a large number of genetic variants. For example, for the first time, we found a significant association between rs9691259 and rs6950179 and PCa risk. Third, we validated the promising associations in the discovery stage in an independent cohort. This replication design likely makes our findings more conclusive.

Despite these strengths, our study still has some limitations. First, our study was conducted retrospectively, which carries an intrinsic bias. Second, our discovery stage study was performed using MassARRAY iPLEX platform, which might represent an incomplete coverage on genetic variants in the IGFBP-3 gene. Although additional imputation analyses may improve the study effectiveness, the possibility that some independent SNPs may be associated with PCa risk could not be excluded. Third, some information was inadequate, especially in the replication stage. For example, PSA levels and Gleason score were missing in many PCa patients. In addition, some other clinical data that might influence PCa risk, such as body mass index and history of diabetes, were not gathered in our database. All this missing information may lead to the deficiency of adjustment of potential confounding factors, as well as further stratification analyses, which might weaken our conclusion to a certain extent. Hence, our study requires further validation from well-designed prospective studies with adequate clinicopathological data.

Conclusion

The current two-stage case-control study found that rs9691259, rs6950179, and rs2854744 in the IGFBP-3 gene were significantly associated with altered PCa risk in the Chinese Han population. Our study adds more evidence to the current knowledge about the relationship between IGFBP-3 polymorphisms and cancer risk, and further highlights the important role of IGFBP-3 in PCa carcinogenesis.

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Disclosure

The authors report no conflicts of interest in this work.

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